Appendix A – Outputs for Estimating MoE

This Appendix presents methods of finding the Test Comparison Value (TCV) in order to have given confidence that the production meets or exceeds the Design Value (DV) for MoE.

In estimating the MoE of a batch, random samples are taken as representative of the batch.

The MoE properties found from testing and analysis of the sampled timber are taken as estimates of the batch properties.

Where the test results are greater than or equal to Test Comparison Values calculated from tables or equations in this Appendix, then the producer has the assigned confidence that the production batch meets or exceeds the Design Value.

The steps in using the appendix are as follows:

- 1. Decide on a level of confidence that the MoE of the material produced meets the Design Value.
- 2. Select the section appropriate for the analysis method (eg non-parametric, lognormal or estimates from average MSG) to be used.
- 3. Find the value of Test Comparison Value (*TCV*):
 - For tables, use the selected Confidence Level, sample size and estimated CoV of the grade to look up value of M, and find $TCV = M \times DV$. Interpolation is possible within the table.
 - For equations, use the required Confidence Level to select a value of *A*, and this is used to find the Test Comparison Value for any value of CoV or sample size.
 - For graphs, use the selected Confidence Level, sample size and estimated CoV of the grade to look up value of M, and find $TCV = M \times DV$. (It is more accurate to use either tables or equations, but the graphs provide a visual representation of trends.
- 4. Use the Test Comparison Value in the monitoring system.

For MoE, outputs have been produced giving information on:

- the mean MoE (Section A.1)
- the 5%ile MoE (Section A.2).

Refer to examples in Section 6 for further guidance.

A.1 Mean MoE

A.1.1 Multipliers (M) for non-parametric mean MoE

$$TCV = M * DV$$

Non Parametric Mean MoE					
CL 95%	CoV				
Sample size	8%	10%	12%	15%	20%
5	1.063	1.080	1.097	1.124	1.173
10	1.044	1.055	1.067	1.085	1.116
20	1.030	1.038	1.046	1.059	1.080
30	1.025	1.031	1.037	1.047	1.064
50	1.019	1.024	1.029	1.036	1.049
100	1.013	1.017	1.020	1.025	1.034
200	1.009	1.012	1.014	1.018	1.024

No	n Paran	netric M	ean MoE		
CL 80%	CoV				
Sample size	8%	10%	12%	15%	20%
5	1.032	1.040	1.048	1.061	1.083
10	1.022	1.028	1.034	1.042	1.057
20	1.016	1.019	1.023	1.029	1.040
30	1.013	1.016	1.019	1.024	1.032
50	1.010	1.012	1.015	1.018	1.025
100	1.007	1.009	1.010	1.013	1.017
200	1.005	1.006	1.007	1.009	1.012

Non Parametric Mean MoE					
CL 90%	CoV				
Sample size	8%	10%	12%	15%	20%
5	1.048	1.061	1.074	1.095	1.130
10	1.034	1.043	1.051	1.065	1.089
20	1.024	1.030	1.036	1.045	1.061
30	1.019	1.024	1.029	1.037	1.049
50	1.015	1.019	1.022	1.028	1.038
100	1.010	1.013	1.016	1.020	1.026
200	1.007	1.009	1.011	1.014	1.019

Non Parametric Mean MoE					
CL 75%	CoV				
Sample size	8%	10%	12%	15%	20%
5	1.025	1.032	1.038	1.048	1.065
10	1.018	1.022	1.027	1.034	1.045
20	1.012	1.016	1.019	1.024	1.032
30	1.010	1.013	1.015	1.019	1.026
50	1.008	1.010	1.012	1.015	1.020
100	1.006	1.007	1.008	1.010	1.014
200	1.004	1.005	1.006	1.007	1.010

No	n Paran	netric M	ean Mo	E	
CL 85%	CoV				
Sample size	8%	10%	12%	15%	20%
5	1.039	1.049	1.059	1.075	1.103
10	1.027	1.034	1.041	1.052	1.071
20	1.019	1.024	1.029	1.036	1.049
30	1.016	1.019	1.023	1.029	1.040
50	1.012	1.015	1.018	1.023	1.030
100	1.008	1.011	1.013	1.016	1.021
200	1.006	1.007	1.009	1.011	1.015

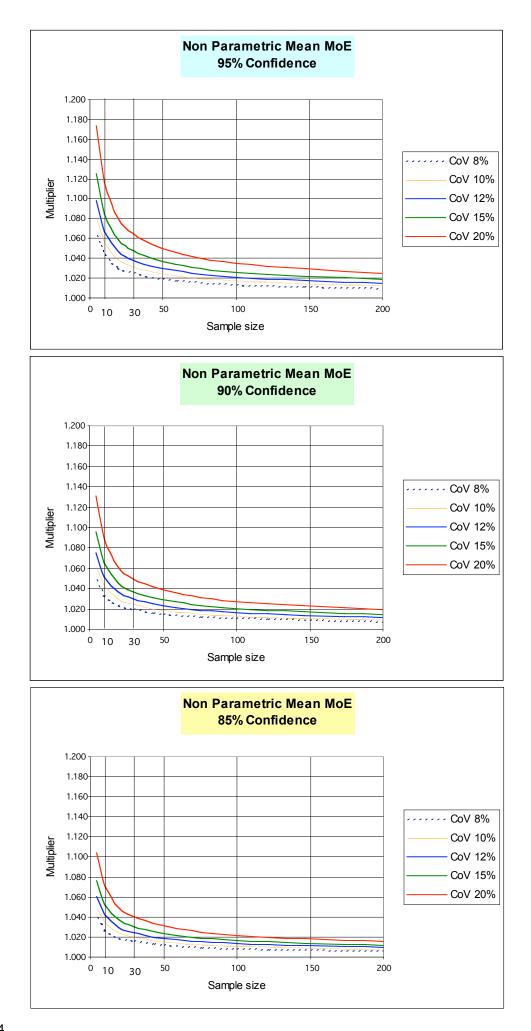
The estimation of mean MoE using non-parametric methods is very simple.

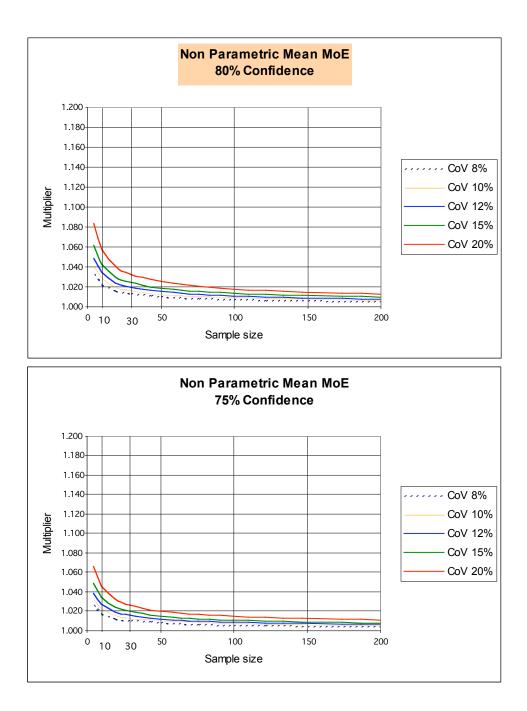
- The estimate is the average of all of the MoE data.
- No distribution is fitted.
- The AVERAGE function in msExcel can be used.

The Test Comparison Value used to accept or reject this estimate is found using the following equation or the preceding Tables.

Equation – Non-parametric mean MoE

	CL	Value of A
	95%	-1.649
$TCV = DV \left \frac{1}{2} \right $	90%	-1.290
CV = DV	85%	-1.045
$1 + A - \overline{\sqrt{x}}$	80%	-0.854
$ \sqrt{n} $	75%	-0.686





A.1.2 Multipliers (M) for Log-normal mean MoE

$$TCV = M * DV$$

Log-Normal Mean MoE					
CL 95%	CoV				
Sample size	8%	10%	12%	15%	20%
5	1.063	1.080	1.098	1.125	1.174
10	1.044	1.055	1.067	1.085	1.117
20	1.031	1.038	1.047	1.059	1.080
30	1.025	1.031	1.038	1.048	1.064
50	1.019	1.024	1.029	1.036	1.049
100	1.013	1.017	1.020	1.025	1.034
200	1.009	1.012	1.014	1.018	1.024

	Log-Nor	mal Mea	n MoE		
CL 80%	CoV				
Sample size	8%	10%	12%	15%	20%
5	1.032	1.040	1.048	1.061	1.083
10	1.022	1.028	1.034	1.043	1.058
20	1.016	1.020	1.024	1.030	1.040
30	1.013	1.016	1.019	1.024	1.032
50	1.010	1.012	1.015	1.019	1.025
100	1.007	1.009	1.010	1.013	1.018
200	1.005	1.006	1.007	1.009	1.012

Log-Normal Mean MoE					
CL 90%	CoV				
Sample size	8%	10%	12%	15%	20%
5	1.049	1.062	1.075	1.095	1.131
10	1.034	1.043	1.052	1.066	1.089
20	1.024	1.030	1.036	1.045	1.062
30	1.019	1.024	1.029	1.037	1.050
50	1.015	1.019	1.023	1.028	1.038
100	1.010	1.013	1.016	1.020	1.027
200	1.007	1.009	1.011	1.014	1.019

	Log-Nor	mal Mea	n MoE		
CL 75%	CoV				
Sample size	8%	10%	12%	15%	20%
5	1.025	1.032	1.039	1.049	1.066
10	1.018	1.022	1.027	1.034	1.046
20	1.013	1.016	1.019	1.024	1.032
30	1.010	1.013	1.015	1.019	1.026
50	1.008	1.010	1.012	1.015	1.020
100	1.006	1.007	1.008	1.011	1.014
200	1.004	1.005	1.006	1.007	1.010

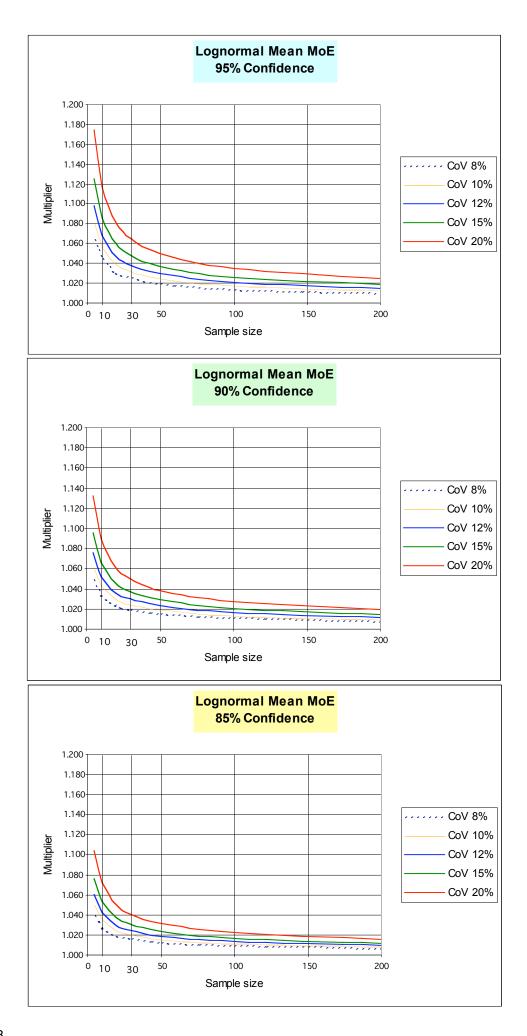
L	og-Nori	nal Mea	an MoE		
CL 85%	CoV				
Sample size	8%	10%	12%	15%	20%
5	1.039	1.049	1.060	1.076	1.104
10	1.027	1.034	1.042	1.053	1.071
20	1.019	1.024	1.029	1.037	1.049
30	1.016	1.020	1.024	1.030	1.040
50	1.012	1.015	1.018	1.023	1.031
100	1.008	1.011	1.013	1.016	1.021
200	1.006	1.007	1.009	1.011	1.015

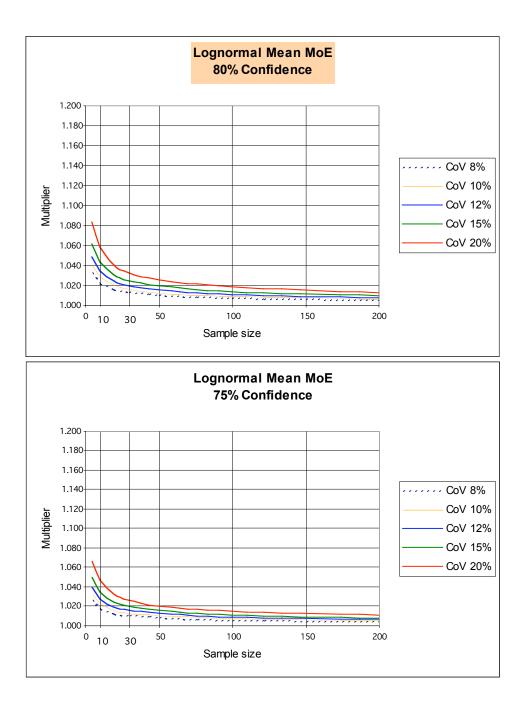
The estimation of mean MoE using a log-normal distribution through the data makes use of the natural logarithms of the raw MoE data (ln(MoE)). An extra column needs to be introduced to an analysis spreadsheet. This same column and many of the calculations necessary for this method can be used to estimate the 5%ile MoE from the log-normal distribution.

- The estimate of the mean MoE is based on the average and the standard deviation of the ln(MoE)data.
- The mean MoE can be found using (*eqn C.9*).
- The LN, EXP, AVERAGE and STDEV functions in msExcel are used.

г л	CL	Value of A
	95%	-1.657
$\left \frac{1}{2}\right $	90%	-1.297
CoV	85%	-1.052
$1 + A - \frac{1}{\Gamma}$	80%	-0.861
\sqrt{n}	75%	-0.693

Equation – Log-normal mean MoE





A.1.3 Multipliers (*M*) for **MSG estimate** of Mean MoE

$$TCV = M * DV$$

MSG Estimate of Mean MoE						
CL 95%		minMSG/avgMSG				
Sample size	65%	70%	75%	80%	85%	
50	1.088	1.072	1.057	1.042	1.028	
100	1.075	1.062	1.048	1.034	1.021	
200	1.067	1.054	1.041	1.029	1.017	
500	1.060	1.048	1.036	1.024	1.013	
1000	1.056	1.044	1.033	1.022	1.011	
2000	1.053	1.042	1.031	1.020	1.009	
5000	1.051	1.040	1.029	1.018	1.008	
10000	1.050	1.039	1.028	1.018	1.007	

MSG Estimate of Mean MoE						
CL 80%		minMSG/avgMSG				
Sample size	65%	70%	75%	80%	85%	
50	1.068	1.054	1.042	1.029	1.017	
100	1.061	1.049	1.037	1.025	1.014	
200	1.057	1.045	1.034	1.022	1.011	
500	1.054	1.042	1.031	1.020	1.009	
1000	1.052	1.040	1.029	1.019	1.008	
2000	1.050	1.039	1.028	1.018	1.007	
5000	1.049	1.038	1.028	1.017	1.007	
10000	1.049	1.038	1.027	1.017	1.006	

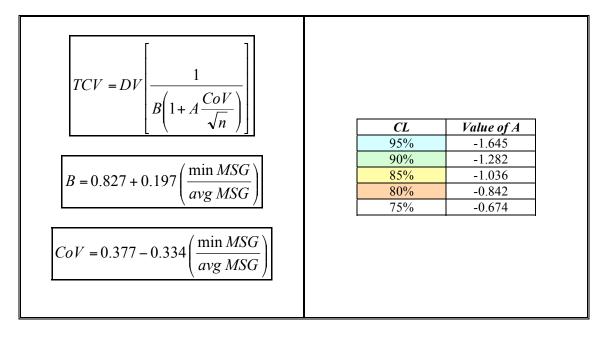
MSG Estimate of Mean MoE						
CL90%		minN	ISG/avg	MSG		
Sample size	65%	70%	75%	80%	85%	
50	1.078	1.064	1.050	1.036	1.023	
100	1.069	1.056	1.043	1.030	1.018	
200	1.063	1.050	1.038	1.026	1.014	
500	1.057	1.045	1.034	1.022	1.011	
1000	1.054	1.043	1.031	1.020	1.010	
2000	1.052	1.041	1.030	1.019	1.008	
5000	1.050	1.039	1.028	1.018	1.007	
10000	1.049	1.038	1.028	1.017	1.007	

MSG Estimate of Mean MoE						
CL75%		minMSG/avgMSG				
Sample size	65%	70%	75%	80%	85%	
50	1.063	1.051	1.039	1.027	1.015	
100	1.059	1.047	1.035	1.023	1.012	
200	1.055	1.044	1.032	1.021	1.010	
500	1.052	1.041	1.030	1.019	1.009	
1000	1.051	1.040	1.029	1.018	1.008	
2000	1.050	1.039	1.028	1.017	1.007	
5000	1.049	1.038	1.027	1.017	1.007	
10000	1.048	1.037	1.027	1.017	1.006	

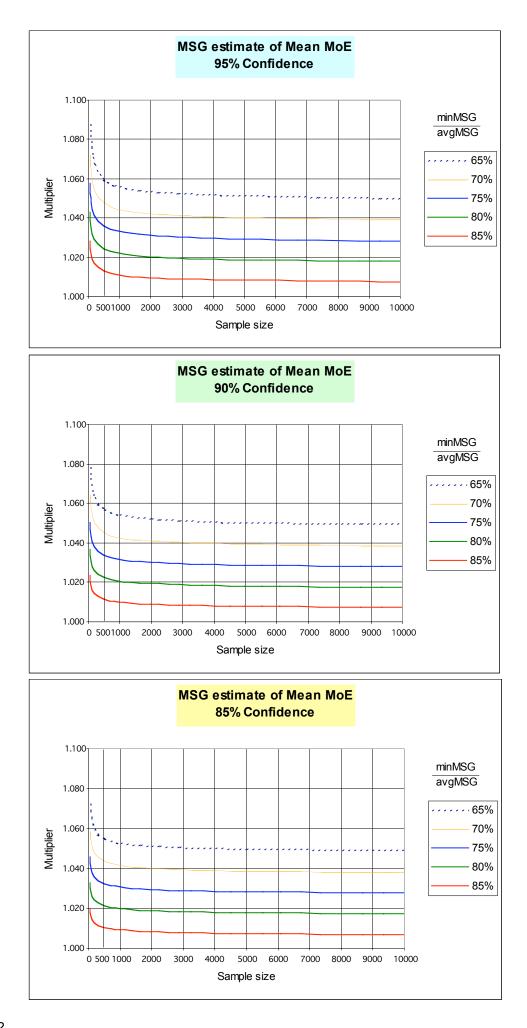
MSG Estimate of Mean MoE						
CL85%		minMSG/avgMSG				
Sample size	65%	70%	75%	80%	85%	
50	1.072	1.059	1.045	1.032	1.020	
100	1.065	1.052	1.040	1.027	1.016	
200	1.060	1.047	1.036	1.024	1.013	
500	1.055	1.043	1.032	1.021	1.010	
1000	1.053	1.041	1.030	1.019	1.009	
2000	1.051	1.040	1.029	1.018	1.008	
5000	1.050	1.039	1.028	1.017	1.007	
10000	1.049	1.038	1.027	1.017	1.007	

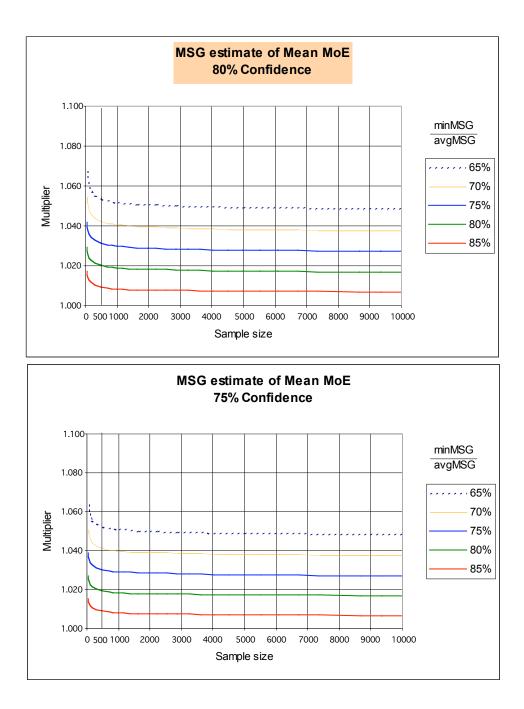
The estimation of mean MoE using the average MSG of each piece in the period of production relies on collection of data already available in most Machine Stress Graders.

- The average MSG of each piece is averaged in the control unit of the MSG and is usually displayed continuously, and is recorded at the end of each run.
- The value recorded is simply compared with the Test Comparison Value.



Equation – MSG estimate of mean MoE





A.2 5%ile MoE

A.2.1 Multipliers (M) for Non-parametric 5%ile MoE

$$TCV = M * DV$$

Non Parametric 5%ile MoE						
CL 95%		CoV				
Sample size	8%	10%	12%	15%	20%	
5	1.155	1.201	1.252	1.336	1.504	
10	1.105	1.134	1.166	1.216	1.311	
20	1.072	1.091	1.112	1.144	1.201	
30	1.058	1.073	1.089	1.114	1.159	
50	1.044	1.056	1.068	1.086	1.119	
100	1.031	1.039	1.047	1.060	1.081	
200	1.022	1.027	1.033	1.041	1.056	

Non Parametric 5%ile MoE						
CL 80%		CoV				
Sample size	8%	10%	12%	15%	20%	
5	1.091	1.116	1.143	1.185	1.263	
10	1.062	1.079	1.097	1.124	1.172	
20	1.043	1.055	1.067	1.085	1.116	
30	1.035	1.044	1.054	1.068	1.093	
50	1.027	1.034	1.041	1.052	1.070	
100	1.019	1.024	1.029	1.036	1.049	
200	1.013	1.017	1.020	1.025	1.034	

Non Parametric 5%ile MoE						
CL 90%		CoV				
Sample size	8%	10%	12%	15%	20%	
5	1.125	1.162	1.200	1.264	1.385	
10	1.085	1.109	1.134	1.173	1.245	
20	1.059	1.075	1.091	1.116	1.162	
30	1.048	1.060	1.073	1.093	1.128	
50	1.036	1.046	1.056	1.071	1.096	
100	1.026	1.032	1.039	1.049	1.066	
200	1.018	1.022	1.027	1.034	1.046	

Non Parametric 5%ile MoE						
CL 75%		CoV				
Sample size	8%	10%	12%	15%	20%	
5	1.078	1.099	1.122	1.157	1.221	
10	1.054	1.068	1.083	1.106	1.147	
20	1.038	1.047	1.057	1.073	1.099	
30	1.030	1.038	1.046	1.059	1.080	
50	1.023	1.029	1.036	1.045	1.061	
100	1.016	1.021	1.025	1.031	1.042	
200	1.012	1.015	1.017	1.022	1.029	

Non Parametric 5%ile MoE						
CL 85%		CoV				
Sample size	8%	10%	12%	15%	20%	
5	1.106	1.136	1.167	1.218	1.314	
10	1.073	1.092	1.113	1.145	1.203	
20	1.050	1.064	1.077	1.098	1.136	
30	1.041	1.051	1.062	1.079	1.108	
50	1.031	1.039	1.048	1.060	1.082	
100	1.022	1.027	1.033	1.042	1.056	
200	1.015	1.019	1.023	1.029	1.039	

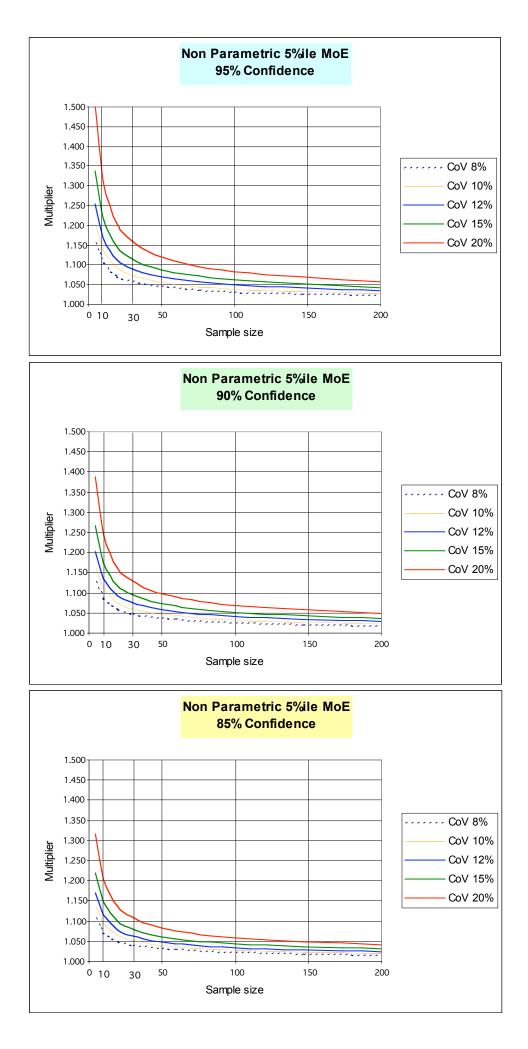
The non-parametric 5% ile can be estimated for test data by ranking the test data, plotting it against its probability and interpolating to find the 5% ile value.

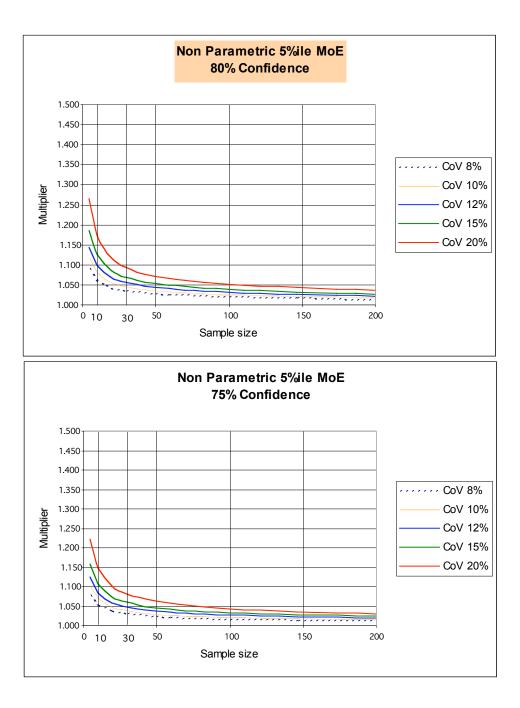
- Detail on the ranking and determination of the 5% ile value is given in Appendix C.2.1.
- The PERCENTILE function in msExcel can be used do this, but has some small errors that may lead to slightly different answers from more rigorously correct analysis methods.

The Test Comparison Value used to accept or reject this estimate is found using the following equation or the preceding Tables.

г	CL	Value of A
	95%	-3.747
$TCV = DV \left \frac{1}{2} \right $	90%	-3.110
$I C V = D V \left[\frac{1}{1 - C O V} \right]$	85%	-2.672
$\left 1 + A - \frac{1}{\sqrt{2}}\right $	80%	-2.325
$\lfloor \sqrt{n} \rfloor$	75%	-2.024

Equation – Non-parametric 5%ile MoE





A.2.2 Multipliers (M) for Log-normal 5%ile MoE

$$TCV = M * DV$$

Log-normal 5%ile MoE						
CL 95%		CoV				
Sample size	8%	10%	12%	15%	20%	
5	1.106	1.136	1.168	1.219	1.316	
10	1.073	1.093	1.113	1.146	1.204	
20	1.050	1.064	1.078	1.099	1.136	
30	1.041	1.051	1.062	1.079	1.109	
50	1.031	1.039	1.048	1.060	1.082	
100	1.022	1.028	1.033	1.042	1.057	
200	1.015	1.019	1.023	1.029	1.039	

	Log-nori	nal 5%il	e MoE							
CL 80%		CoV								
Sample size	8%	8% 10% 12% 15% 20%								
5	1.054	1.069	1.084	1.107	1.148					
10	1.038	1.048	1.058	1.073	1.100					
20	1.026	1.033	1.040	1.051	1.069					
30	1.021	1.027	1.033	1.041	1.055					
50	1.017	1.021	1.025	1.031	1.042					
100	1.012	1.015	1.018	1.022	1.030					
200	1.008	1.010	1.012	1.015	1.021					

L	Log-normal 5%ile MoE											
CL 90%			CoV									
Sample size	8%	8% 10% 12% 15% 20%										
5	1.082	1.105	1.128	1.166	1.234							
10	1.057	1.072	1.088	1.112	1.155							
20	1.039	1.050	1.060	1.077	1.105							
30	1.032	1.040	1.049	1.062	1.084							
50	1.025	1.031	1.037	1.047	1.064							
100	1.017	1.022	1.026	1.033	1.044							
200	1.012	1.015	1.018	1.023	1.031							

	Log-normal 5%ile MoE											
CL 75%	CoV											
Sample size	8%	10%	12%	15%	20%							
5	1.044	1.056	1.067	1.086	1.118							
10	1.031	1.039	1.047	1.059	1.081							
20	1.022	1.027	1.033	1.041	1.056							
30	1.018	1.022	1.026	1.033	1.045							
50	1.014	1.017	1.020	1.026	1.034							
100	1.010	1.012	1.014	1.018	1.024							
200	1.007	1.008	1.010	1.013	1.017							

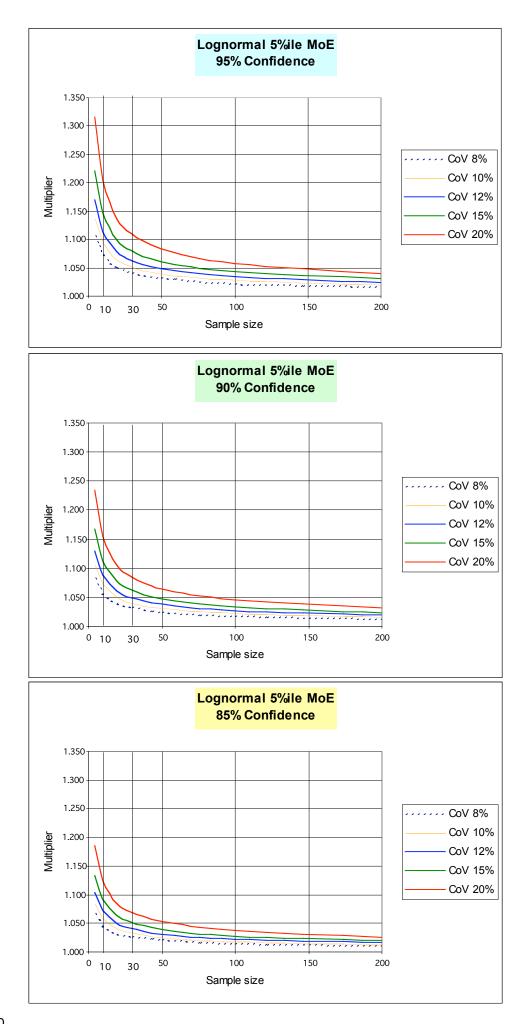
L	Log-normal 5%ile MoE											
CL 85%		CoV										
Sample size	8%	8% 10% 12% 15% 20%										
5	1.066	1.084	1.103	1.132	1.184							
10	1.046	1.058	1.071	1.090	1.123							
20	1.032	1.040	1.049	1.062	1.084							
30	1.026	1.033	1.040	1.050	1.068							
50	1.020	1.025	1.030	1.038	1.052							
100	1.014	1.018	1.021	1.027	1.036							
200	1.010	1.012	1.015	1.019	1.025							

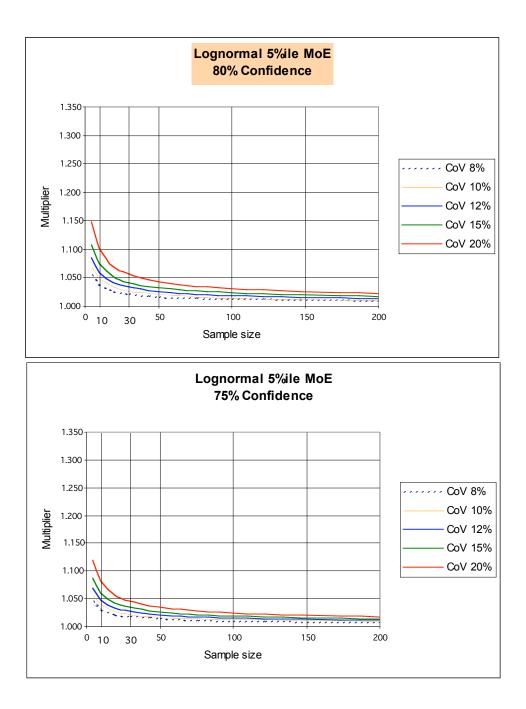
The estimation of 5% ile MoE using a log-normal distribution through the data makes use of the natural logarithms of the raw MoE data (ln(MoE)). An extra column needs to be introduced to an analysis spreadsheet. This same column and many of the calculations necessary for this method can also be used to estimate the mean MoE from the log-normal distribution.

- The estimate of the 5% ile MoE is based on the average and the standard deviation of the ln(MoE) data.
- The 5%ile MoE can be found using (*eqn C.11*).
- The LN, EXP, AVERAGE and STDEV functions in msExcel are used.

 $TCV = DV \left[\frac{1}{1 + A \frac{CoV}{\sqrt{n}}} \right]$ $\frac{CL}{85\%} \frac{Value of A}{-2.682}$ 90%
-2.121
85%
-1.737
80%
-1.438
75%
-1.178

Equation – Log-normal 5%ile MoE





Appendix B – Outputs for Estimating 5%ile Strength

This Appendix presents methods of finding the Test Comparison Value (TCV) in order to have given confidence that the production meets or exceeds the Design Value (DV).

In estimating the strength (either bending or tension) of all of the pieces in a batch, random samples are taken as representative of the batch.

The strength properties found from testing and analysis of the sampled timber are taken as estimates of the batch properties.

Where the test results are greater than or equal to Test Comparison Values calculated from tables or equations in this Appendix, then the producer has the assigned confidence that the production batch meets or exceeds the Design Value.

The steps in using the appendix are as follows:

- 1. Decide on a level of confidence that the strength of the material produced meets the Design Value.
- 2. Select the section appropriate for the analysis method (eg non-parametric, lognormal, log-normal tail fit, 2 parameter Weibull tail fit) to be used.
- 3. Find the value of Test Comparison Value (*TCV*):
 - For tables, use the selected Confidence Level, sample size and estimated CoV of the grade to look up value of M, and find $TCV = M \times DV$. Interpolation is possible within the table.
 - For equations, use the required Confidence Level to select a value of *A*, and this is used to find the Test Comparison Value for any value of CoV or sample size.
 - For graphs, use the selected Confidence Level, sample size and estimated CoV of the grade to look up value of M, and find $TCV = M \times DV$. (It is more accurate to use either tables or equations, but the graphs present a visual representation of trends.
- 4. Use the Test Comparison Value in the monitoring system.

Refer to examples in Section 6 for further guidance.

B.1 Multipliers (M) for Non-parametric 5%ile Strength

TCV = M * DV

	Non-parametric 5%ile Strength										
CL 95%				CoV	· (%)						
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%			
5	1.090	1.198	1.330	1.494	1.705	1.985	2.374	2.954			
10	1.062	1.132	1.213	1.305	1.413	1.540	1.693	1.879			
20	1.043	1.090	1.142	1.198	1.261	1.330	1.407	1.494			
30	1.035	1.072	1.113	1.156	1.203	1.254	1.309	1.370			
50	1.027	1.055	1.085	1.117	1.150	1.186	1.224	1.265			
100	1.019	1.038	1.059	1.080	1.102	1.125	1.149	1.174			
200	1.013	1.027	1.041	1.055	1.070	1.085	1.101	1.117			

	Non-parametric 5%ile Strength									
CL 90%				CoV	′ (%)					
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%		
5	1.074	1.159	1.260	1.379	1.523	1.701	1.926	2.220		
10	1.051	1.108	1.171	1.241	1.321	1.411	1.515	1.636		
20	1.036	1.074	1.115	1.159	1.207	1.260	1.317	1.379		
30	1.029	1.059	1.092	1.126	1.163	1.202	1.244	1.289		
50	1.022	1.045	1.070	1.095	1.122	1.150	1.179	1.210		
100	1.016	1.032	1.048	1.065	1.083	1.102	1.120	1.140		
200	1.011	1.022	1.034	1.045	1.057	1.070	1.082	1.095		

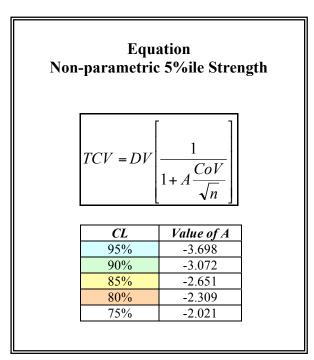
	Non-parametric 5%ile Strength									
CL 85%				CoV	′ (%)					
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%		
5	1.063	1.135	1.216	1.311	1.421	1.552	1.709	1.902		
10	1.044	1.092	1.144	1.201	1.265	1.336	1.415	1.505		
20	1.031	1.063	1.098	1.135	1.174	1.216	1.262	1.311		
30	1.025	1.051	1.078	1.107	1.138	1.170	1.204	1.240		
50	1.019	1.039	1.060	1.081	1.103	1.127	1.151	1.176		
100	1.013	1.027	1.041	1.056	1.071	1.086	1.102	1.119		
200	1.009	1.019	1.029	1.039	1.049	1.060	1.070	1.081		

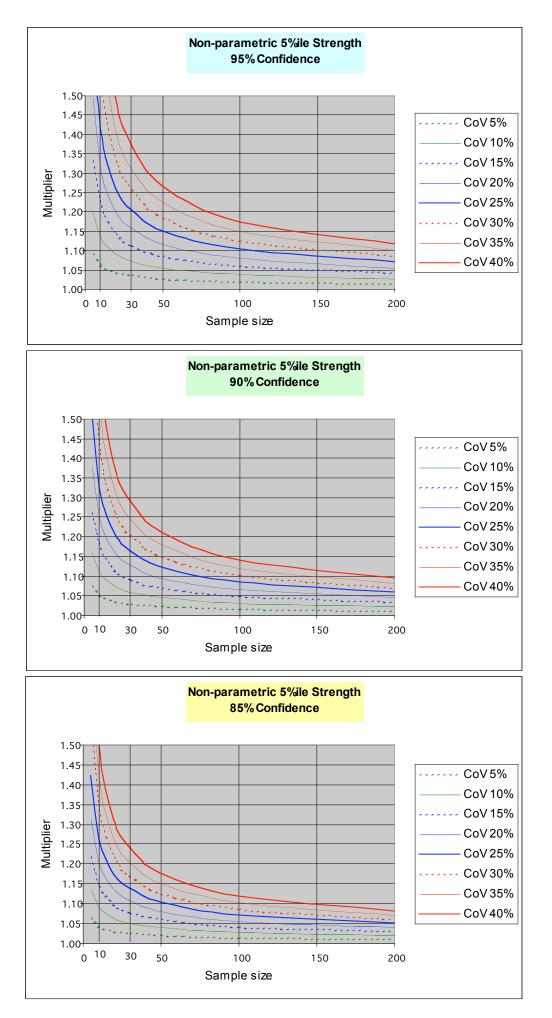
	Non-parametric 5%ile Strength											
CL 80%				CoV	′ (%)							
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%				
5	1.054	1.115	1.183	1.260	1.348	1.449	1.566	1.704				
10	1.038	1.079	1.123	1.171	1.223	1.281	1.343	1.413				
20	1.027	1.054	1.084	1.115	1.148	1.183	1.221	1.260				
30	1.022	1.044	1.068	1.092	1.118	1.145	1.173	1.203				
50	1.017	1.034	1.052	1.070	1.089	1.109	1.129	1.150				
100	1.012	1.024	1.036	1.048	1.061	1.074	1.088	1.102				
200	1.008	1.017	1.025	1.034	1.043	1.052	1.061	1.070				

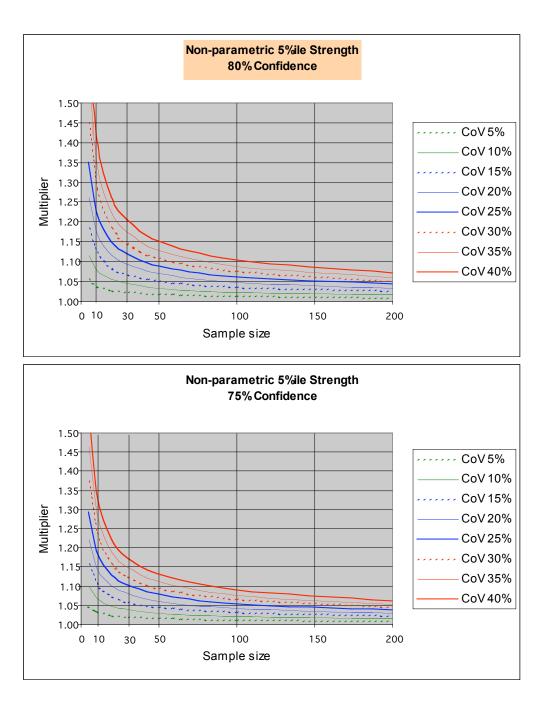
	Non-parametric 5%ile Strength										
CL 75%		CoV (%)									
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%			
5	1.047	1.099	1.157	1.221	1.292	1.372	1.463	1.566			
10	1.033	1.068	1.106	1.147	1.190	1.237	1.288	1.343			
20	1.023	1.047	1.073	1.099	1.127	1.157	1.188	1.221			
30	1.019	1.038	1.059	1.080	1.102	1.124	1.148	1.173			
50	1.014	1.029	1.045	1.061	1.077	1.094	1.111	1.129			
100	1.010	1.021	1.031	1.042	1.053	1.065	1.076	1.088			
200	1.007	1.014	1.022	1.029	1.037	1.045	1.053	1.061			

The non-parametric 5% ile can be estimated for test data by ranking the test data, plotting it against its probability and interpolating to find the 5% ile value.

- Detail on the ranking and determination of the 5% ile value is given in Appendix C.2.1.
- The PERCENTILE function in msExcel can be used do this, but has some small errors that may lead to slightly different answers from more rigorously correct analysis methods.







B.2 Multipliers (M) for Log-normal 5%ile Strength (all data)

$$TCV = M * DV$$

	Log-normal 5%ile Strength (All data)										
CL 95%		-		CoV	(%)						
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%			
5	1.063	1.135	1.217	1.312	1.423	1.554	1.713	1.907			
10	1.044	1.092	1.144	1.202	1.266	1.337	1.417	1.507			
20	1.031	1.063	1.098	1.135	1.175	1.217	1.263	1.312			
30	1.025	1.051	1.079	1.108	1.138	1.170	1.205	1.241			
50	1.019	1.039	1.060	1.081	1.104	1.127	1.152	1.177			
100	1.013	1.027	1.042	1.056	1.071	1.087	1.103	1.119			
200	1.009	1.019	1.029	1.039	1.049	1.060	1.070	1.081			

	Log-normal 5%ile Strength (All data)										
CL 90%				CoV	′ (%)						
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%			
5	1.049	1.104	1.164	1.232	1.308	1.393	1.491	1.603			
10	1.034	1.071	1.111	1.153	1.200	1.249	1.304	1.363			
20	1.024	1.049	1.076	1.104	1.133	1.164	1.197	1.232			
30	1.020	1.040	1.061	1.083	1.106	1.130	1.155	1.182			
50	1.015	1.031	1.047	1.063	1.080	1.098	1.116	1.135			
100	1.011	1.021	1.033	1.044	1.056	1.067	1.079	1.092			
200	1.007	1.015	1.023	1.031	1.039	1.047	1.055	1.063			

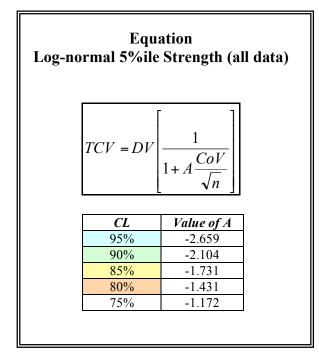
	Log-normal 5%ile Strength (All data)									
CL 85%				CoV	′ (%)					
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%		
5	1.040	1.084	1.131	1.183	1.240	1.303	1.372	1.449		
10	1.028	1.058	1.089	1.123	1.159	1.196	1.237	1.280		
20	1.020	1.040	1.062	1.084	1.107	1.131	1.157	1.183		
30	1.016	1.033	1.050	1.067	1.086	1.105	1.124	1.145		
50	1.012	1.025	1.038	1.051	1.065	1.079	1.094	1.109		
100	1.009	1.018	1.027	1.036	1.045	1.055	1.064	1.074		
200	1.006	1.012	1.019	1.025	1.032	1.038	1.045	1.051		

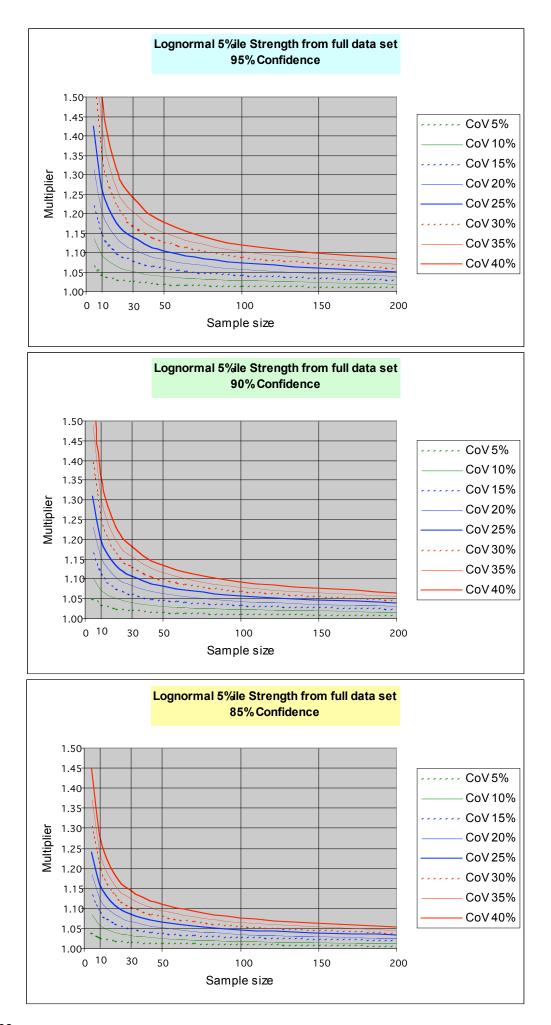
Log-normal 5%ile Strength (All data)										
CL 80%		CoV (%)								
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%		
5	1.033	1.068	1.106	1.147	1.190	1.238	1.289	1.344		
10	1.023	1.047	1.073	1.099	1.128	1.157	1.188	1.221		
20	1.016	1.033	1.050	1.068	1.087	1.106	1.126	1.147		
30	1.013	1.027	1.041	1.055	1.070	1.085	1.101	1.117		
50	1.010	1.021	1.031	1.042	1.053	1.065	1.076	1.088		
100	1.007	1.015	1.022	1.029	1.037	1.045	1.053	1.061		
200	1.005	1.010	1.015	1.021	1.026	1.031	1.037	1.042		

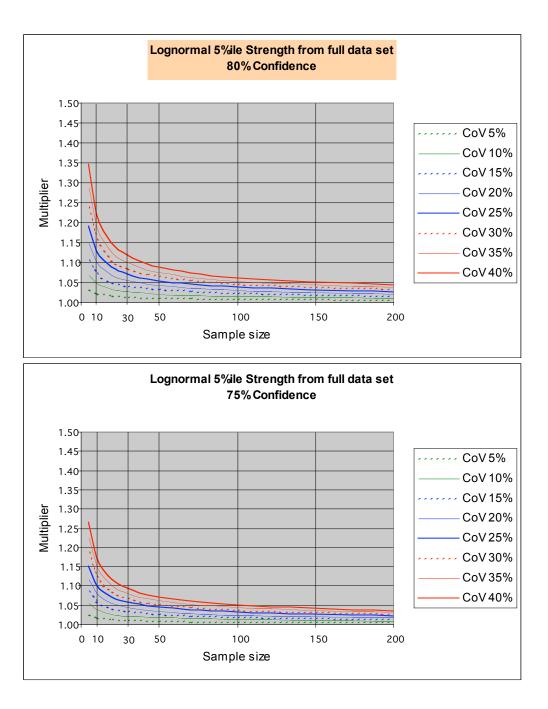
	Log-normal 5%ile Strength (All data)										
CL 75%		CoV (%)									
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%			
5	1.027	1.055	1.085	1.117	1.151	1.187	1.225	1.265			
10	1.019	1.038	1.059	1.080	1.102	1.125	1.149	1.174			
20	1.013	1.027	1.041	1.055	1.070	1.085	1.101	1.117			
30	1.011	1.022	1.033	1.045	1.057	1.069	1.081	1.094			
50	1.008	1.017	1.025	1.034	1.043	1.052	1.062	1.071			
100	1.006	1.012	1.018	1.024	1.030	1.036	1.043	1.049			
200	1.004	1.008	1.013	1.017	1.021	1.025	1.030	1.034			

The estimation of 5%ile strength using a log-normal distribution through the data makes use of the natural logarithms of the raw strength data (ln(strength)). An extra column needs to be introduced to an analysis spreadsheet.

- The estimate of the 5% ile strength is based on the average and the standard deviation of the ln(strength)data.
- The 5% ile strength can be found using (*eqn C.11*).
- The LN, EXP AVERAGE and STDEV functions in msExcel are used.







B.3 Multipliers (M) for Log-normal 5%ile Strength from mean (Tight CoV)

Tables give value	5 01 <i>IVI</i>							
		Log-norn	nal 5%ile St	rength (Tigł	nt CoV)			
CL 95%		_	_	CoV	(%)		_	
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%
5	1.051	1.107	1.170	1.240	1.319	1.410	1.513	1.632
10	1.035	1.074	1.114	1.159	1.207	1.259	1.315	1.377
20	1.025	1.051	1.078	1.107	1.138	1.170	1.204	1.240
30	1.020	1.041	1.063	1.086	1.110	1.135	1.161	1.188

1.048

1.034

1.024

1.065

1.045

1.032

1.083

1.057

1.040

1.101

1.069

1.048

1.120

1.082

1.057

1.140

1.095

1.065

TCV = M * DV

Tables give values of M

50

100

200

1.016

1.011

1.008

1.032

1.022

1.016

	Log-normal 5%ile Strength (Tight CoV)										
CL 90%		CoV (%)									
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%			
5	1.042	1.088	1.138	1.193	1.253	1.320	1.394	1.477			
10	1.029	1.061	1.094	1.129	1.167	1.207	1.250	1.296			
20	1.021	1.042	1.064	1.088	1.112	1.138	1.165	1.193			
30	1.017	1.034	1.052	1.071	1.090	1.110	1.130	1.152			
50	1.013	1.026	1.040	1.054	1.068	1.083	1.098	1.114			
100	1.009	1.018	1.028	1.037	1.047	1.057	1.067	1.078			
200	1.006	1.013	1.020	1.026	1.033	1.040	1.047	1.054			

	Log-normal 5%ile Strength (Tight CoV)										
CL 85%		CoV (%)									
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%			
5	1.036	1.075	1.117	1.163	1.212	1.266	1.324	1.389			
10	1.025	1.052	1.080	1.110	1.141	1.174	1.209	1.247			
20	1.018	1.036	1.055	1.075	1.096	1.117	1.140	1.163			
30	1.014	1.029	1.045	1.061	1.077	1.094	1.111	1.129			
50	1.011	1.023	1.034	1.046	1.059	1.071	1.084	1.097			
100	1.008	1.016	1.024	1.032	1.041	1.049	1.058	1.067			
200	1.006	1.011	1.017	1.023	1.028	1.034	1.040	1.046			

Log-normal 5%ile Strength (Tight CoV)										
CL 80%		CoV (%)								
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%		
5	1.032	1.065	1.101	1.140	1.181	1.226	1.273	1.325		
10	1.022	1.045	1.070	1.095	1.122	1.150	1.179	1.210		
20	1.016	1.032	1.048	1.065	1.083	1.101	1.120	1.140		
30	1.013	1.026	1.039	1.053	1.067	1.081	1.096	1.111		
50	1.010	1.020	1.030	1.040	1.051	1.062	1.073	1.084		
100	1.007	1.014	1.021	1.028	1.036	1.043	1.050	1.058		
200	1.005	1.010	1.015	1.020	1.025	1.030	1.035	1.040		

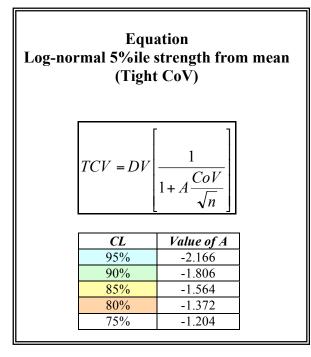
	Log-normal 5%ile Strength (Tight CoV)										
CL 75%		CoV (%)									
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%			
5	1.028	1.057	1.088	1.121	1.156	1.193	1.232	1.275			
10	1.019	1.040	1.061	1.082	1.105	1.129	1.154	1.180			
20	1.014	1.028	1.042	1.057	1.072	1.088	1.104	1.121			
30	1.011	1.022	1.034	1.046	1.058	1.071	1.083	1.096			
50	1.009	1.017	1.026	1.035	1.044	1.054	1.063	1.073			
100	1.006	1.012	1.018	1.025	1.031	1.037	1.044	1.051			
200	1.004	1.009	1.013	1.017	1.022	1.026	1.031	1.035			

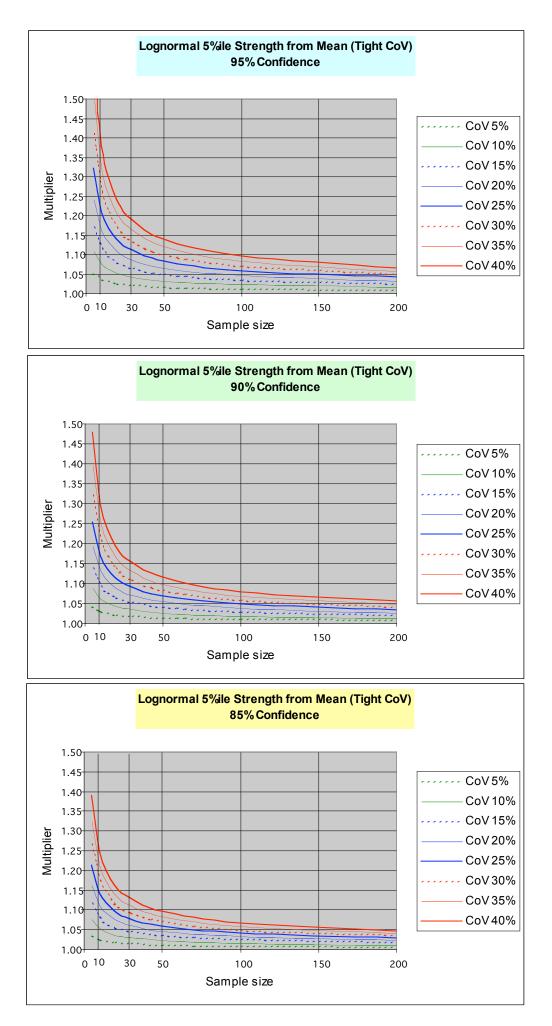
The estimation of 5% ile strength from the mean using a log-normal distribution through the data makes use of the natural logarithms of the raw strength data (ln(strength)). An extra column needs to be introduced to an analysis spreadsheet.

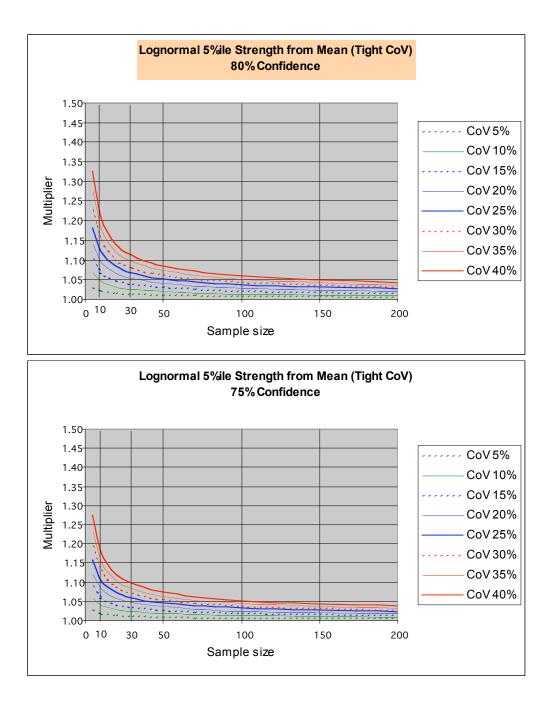
- The estimate of the 5% ile strength is based on the average of the ln(strength) data.
- *(Eqn C.12)* is used to find the standard deviation to be used in the analysis directly from the long-term CoV of the grade.
- The 5% ile strength can be found using (*eqn C.11*).
- The LN EXP and AVERAGE functions in msExcel are used.

The Test Comparison Value used to accept or reject this estimate is found using the following equation or the preceding Tables.

This equation is used when the estimated CoV of production is within 5% of the long-term CoV (ie. Between 0.95 CoV and 1.05 CoV).







B.4 Multipliers for Log-normal 5%ile Strength from mean (Loose CoV)

$$TCV = M * DV$$

	Log-normal 5%ile Strength (Loose CoV)								
CL 95%				CoV	′ (%)				
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%	
5	1.064	1.137	1.220	1.317	1.430	1.565	1.728	1.928	
10	1.044	1.093	1.146	1.205	1.270	1.343	1.424	1.516	
20	1.031	1.064	1.099	1.137	1.177	1.220	1.267	1.317	
30	1.025	1.052	1.080	1.109	1.140	1.173	1.208	1.245	
50	1.019	1.040	1.061	1.082	1.105	1.129	1.154	1.180	
100	1.014	1.028	1.042	1.057	1.072	1.088	1.104	1.121	
200	1.010	1.019	1.029	1.040	1.050	1.061	1.071	1.082	

	Log-normal 5%ile Strength (Loose CoV)								
CL 90%				CoV	′ (%)				
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%	
5	1.055	1.116	1.185	1.263	1.352	1.455	1.574	1.715	
10	1.038	1.080	1.124	1.173	1.226	1.284	1.348	1.418	
20	1.027	1.055	1.085	1.116	1.150	1.185	1.223	1.263	
30	1.022	1.044	1.068	1.093	1.119	1.146	1.175	1.205	
50	1.017	1.034	1.052	1.071	1.090	1.110	1.130	1.152	
100	1.012	1.024	1.036	1.049	1.062	1.075	1.089	1.103	
200	1.008	1.017	1.025	1.034	1.043	1.052	1.061	1.071	

	Log-normal 5%ile Strength (Loose CoV)								
CL 85%		CoV (%)							
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%	
5	1.049	1.103	1.163	1.230	1.305	1.389	1.486	1.596	
10	1.034	1.071	1.110	1.152	1.198	1.247	1.301	1.359	
20	1.024	1.049	1.075	1.103	1.132	1.163	1.195	1.230	
30	1.019	1.040	1.061	1.083	1.105	1.129	1.154	1.180	
50	1.015	1.030	1.046	1.063	1.080	1.097	1.115	1.134	
100	1.011	1.021	1.032	1.044	1.055	1.067	1.079	1.091	
200	1.007	1.015	1.023	1.030	1.038	1.046	1.055	1.063	

	Log-normal 5%ile Strength (Loose CoV)								
CL 80%				CoV	′ (%)				
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%	
5	1.044	1.093	1.146	1.204	1.269	1.341	1.422	1.513	
10	1.031	1.064	1.099	1.136	1.176	1.219	1.266	1.315	
20	1.022	1.044	1.068	1.093	1.119	1.146	1.174	1.204	
30	1.018	1.036	1.055	1.074	1.095	1.116	1.138	1.161	
50	1.014	1.028	1.042	1.057	1.072	1.087	1.104	1.120	
100	1.010	1.019	1.029	1.039	1.050	1.060	1.071	1.082	
200	1.007	1.014	1.021	1.028	1.035	1.042	1.049	1.057	

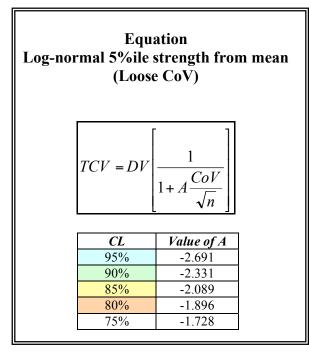
		Log-norm	al 5%ile Str	ength (Loos	se CoV)				
CL 75%		CoV (%)							
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%	
5	1.040	1.084	1.131	1.183	1.239	1.302	1.371	1.447	
10	1.028	1.058	1.089	1.123	1.158	1.196	1.236	1.280	
20	1.020	1.040	1.062	1.084	1.107	1.131	1.156	1.183	
30	1.016	1.033	1.050	1.067	1.086	1.105	1.124	1.144	
50	1.012	1.025	1.038	1.051	1.065	1.079	1.094	1.108	
100	1.009	1.018	1.027	1.036	1.045	1.055	1.064	1.074	
200	1.006	1.012	1.019	1.025	1.032	1.038	1.045	1.051	

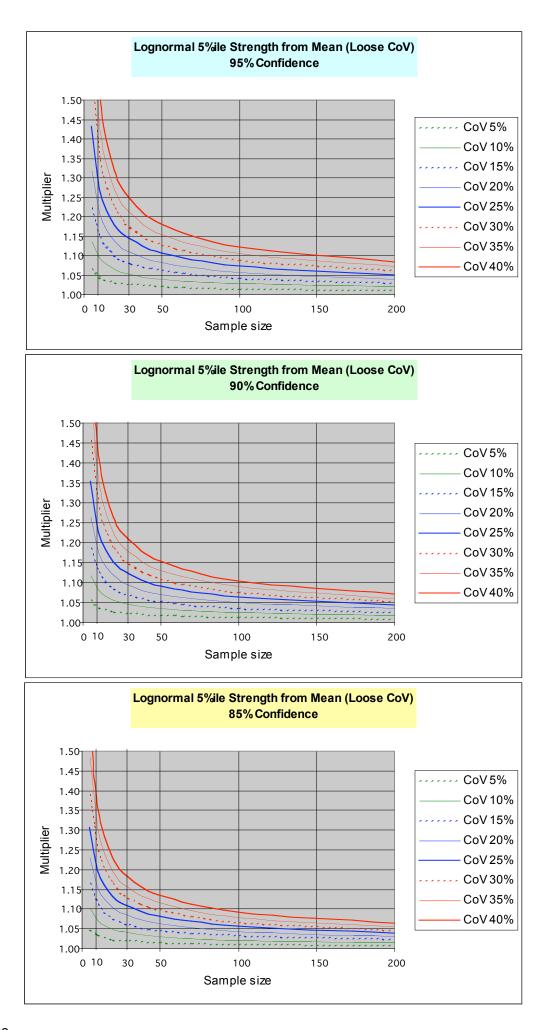
The estimation of 5% ile strength from the mean using a log-normal distribution through the data makes use of the natural logarithms of the raw strength data (ln(strength)). An extra column needs to be introduced to an analysis spreadsheet.

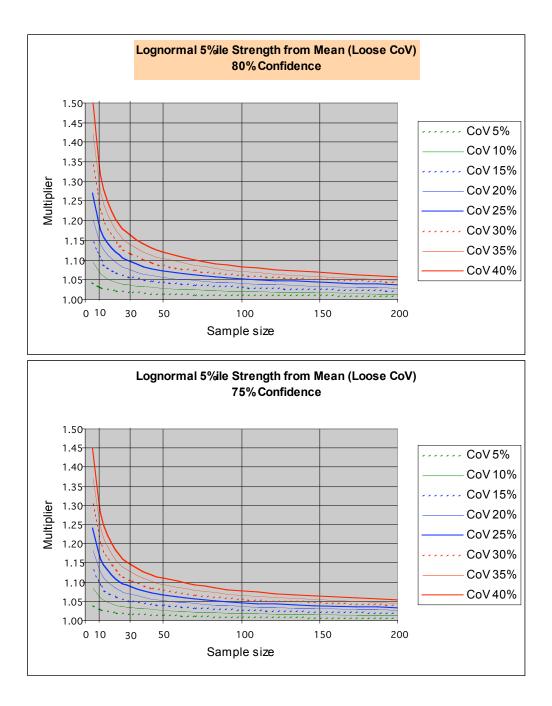
- The estimate of the 5% ile strength is based on the average of the ln(strength) data.
- *(Eqn C.12)* is used to find the standard deviation to be used in the analysis directly from the long-term CoV of the grade.
- The 5% ile strength can be found using (*eqn C.11*).
- The LN,EXP and AVERAGE functions in msExcel are used.

The Test Comparison Value used to accept or reject this estimate is found using the following equation or the preceding Tables.

This equation is used when the estimated CoV of production is within 10% of the long-term CoV (ie. Between 0.90 CoV and 1.10 CoV).







B.5 Multipliers for Log-normal 5%ile Strength (Tail fit)

$$TCV = M * DV$$

	Log-normal 5%ile Strength (Tail fit)							
CL 95%		CoV (%)						
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%
5	1.071	1.154	1.250	1.363	1.499	1.665	1.873	2.139
10	1.049	1.104	1.164	1.232	1.308	1.394	1.491	1.604
20	1.034	1.071	1.111	1.154	1.200	1.250	1.304	1.363
30	1.028	1.057	1.089	1.122	1.157	1.195	1.235	1.278
50	1.022	1.044	1.067	1.092	1.118	1.145	1.173	1.203
100	1.015	1.031	1.047	1.063	1.080	1.098	1.116	1.135
200	1.011	1.022	1.033	1.044	1.056	1.067	1.080	1.092

	Log-normal 5%ile Strength (Tail fit)							
CL 90%				CoV	′ (%)			
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%
5	1.056	1.119	1.190	1.271	1.364	1.470	1.596	1.744
10	1.039	1.082	1.128	1.178	1.232	1.292	1.359	1.432
20	1.027	1.056	1.087	1.119	1.154	1.190	1.229	1.271
30	1.022	1.046	1.070	1.095	1.122	1.150	1.180	1.211
50	1.017	1.035	1.053	1.072	1.092	1.113	1.134	1.156
100	1.012	1.024	1.037	1.050	1.063	1.077	1.091	1.105
200	1.009	1.017	1.026	1.035	1.044	1.053	1.063	1.072

	Log pormal 5% its Strongth (Toil fit)									
	Log-normal 5%ile Strength (Tail fit)									
CL 85%		CoV (%)								
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%		
5	1.046	1.097	1.153	1.215	1.284	1.362	1.449	1.549		
10	1.032	1.067	1.104	1.143	1.186	1.231	1.281	1.334		
20	1.023	1.046	1.071	1.097	1.124	1.153	1.183	1.215		
30	1.018	1.038	1.057	1.078	1.099	1.122	1.145	1.169		
50	1.014	1.029	1.044	1.059	1.075	1.092	1.109	1.126		
100	1.010	1.020	1.031	1.041	1.052	1.063	1.074	1.086		
200	1.007	1.014	1.021	1.029	1.036	1.044	1.052	1.059		

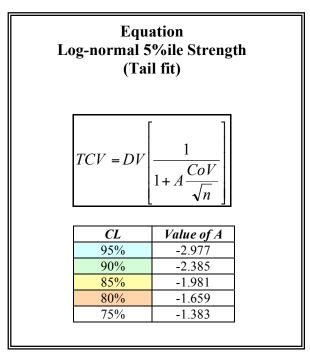
	Log-normal 5%ile Strength (Tail fit)								
CL 80%				CoV	′ (%)				
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%	
5	1.039	1.080	1.125	1.174	1.228	1.286	1.351	1.422	
10	1.027	1.055	1.085	1.117	1.151	1.187	1.225	1.266	
20	1.019	1.039	1.059	1.080	1.102	1.125	1.149	1.174	
30	1.015	1.031	1.048	1.064	1.082	1.100	1.119	1.138	
50	1.012	1.024	1.036	1.049	1.062	1.076	1.089	1.104	
100	1.008	1.017	1.026	1.034	1.043	1.052	1.062	1.071	
200	1.006	1.012	1.018	1.024	1.030	1.036	1.043	1.049	

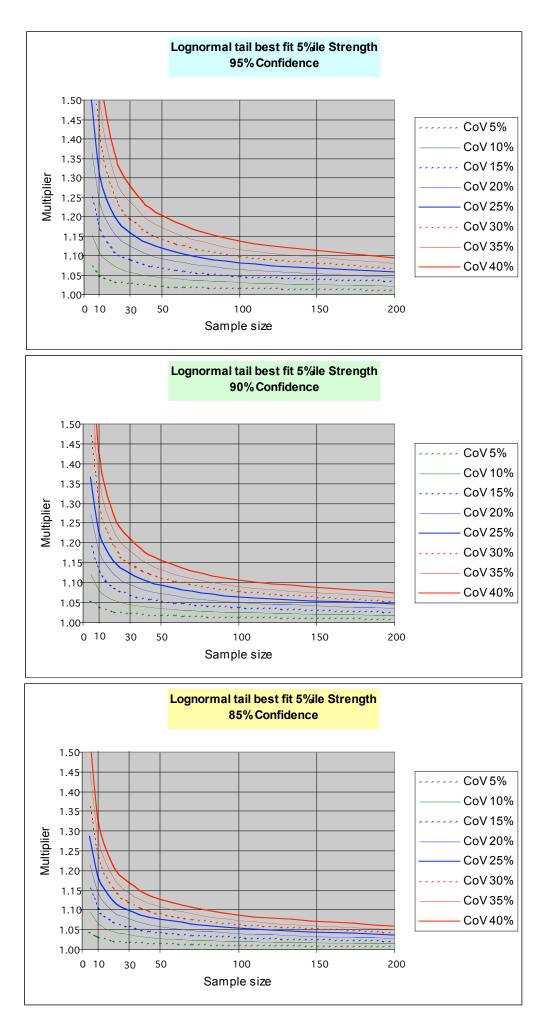
		Log-nc	ormal 5%ile	Strength (Ta	ail fit)			
CL 75%		CoV (%)						
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%
5	1.032	1.066	1.102	1.141	1.183	1.228	1.276	1.329
10	1.022	1.046	1.070	1.096	1.123	1.151	1.181	1.212
20	1.016	1.032	1.049	1.066	1.084	1.102	1.121	1.141
30	1.013	1.026	1.039	1.053	1.067	1.082	1.097	1.112
50	1.010	1.020	1.030	1.041	1.051	1.062	1.073	1.085
100	1.007	1.014	1.021	1.028	1.036	1.043	1.051	1.059
200	1.005	1.010	1.015	1.020	1.025	1.030	1.035	1.041

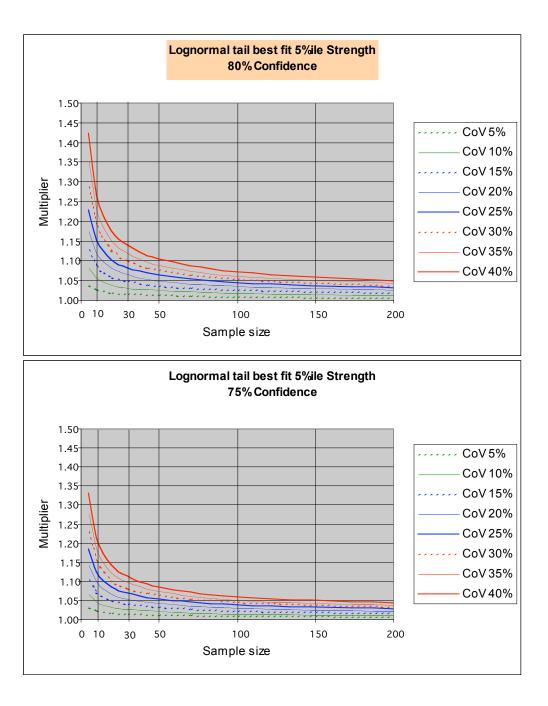
The estimation of 5% ile strength using a log-normal distribution through the tail of the data makes use of the natural logarithms of the raw strength data (ln(strength)). Extra columns need to be introduced to an analysis spreadsheet, to accommodate ln(strength) and probability.

- The data is ranked and probabilities (*pr*) determined in accordance with Appendix C.2.4.
- Ln(strength) is plotted against z(pr) and the mean and standard deviation of the log-normal distribution are given by the intercept and the slope of the line respectively.
- The 5% ile strength can be found using (*eqn C.11*).
- The LN, EXP, SLOPE and INTERCEPT functions in msExcel are used.

The Test Comparison Value used to accept or reject this estimate is found using the following equation or the preceding Tables.







B.6 Multipliers for Log-normal 5%ile Strength from mean (Tail fit, Tight CoV)

TCV = M	*DV
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	Log-normal 5%ile Strength (Tail fit, Tight CoV)								
CL 95%		CoV (%)							
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%	
5	1.057	1.122	1.194	1.277	1.372	1.482	1.611	1.765	
10	1.040	1.083	1.130	1.181	1.237	1.298	1.366	1.442	
20	1.028	1.057	1.088	1.122	1.157	1.194	1.234	1.277	
30	1.023	1.046	1.071	1.097	1.124	1.153	1.183	1.215	
50	1.017	1.035	1.054	1.074	1.094	1.115	1.136	1.159	
100	1.012	1.025	1.038	1.051	1.064	1.078	1.093	1.107	
200	1.009	1.017	1.026	1.035	1.045	1.054	1.064	1.074	

	Log-normal 5%ile Strength (Tail fit, Tight CoV)										
CL 90%		CoV (%)									
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%			
5	1.047	1.098	1.155	1.218	1.288	1.366	1.455	1.556			
10	1.033	1.067	1.105	1.145	1.188	1.234	1.284	1.338			
20	1.023	1.047	1.072	1.098	1.126	1.155	1.185	1.218			
30	1.019	1.038	1.058	1.079	1.100	1.123	1.146	1.171			
50	1.014	1.029	1.044	1.060	1.076	1.093	1.110	1.127			
100	1.010	1.020	1.031	1.042	1.053	1.064	1.075	1.087			
200	1.007	1.014	1.022	1.029	1.037	1.044	1.052	1.060			

	Log-normal 5%ile Strength (Tail fit, Tight CoV)										
CL 85%		CoV (%)									
No. Sample	5%	5% 10% 15% 20% 25% 30% 35% 40%									
5	1.040	1.083	1.130	1.181	1.237	1.298	1.366	1.442			
10	1.028	1.057	1.088	1.122	1.157	1.194	1.234	1.277			
20	1.020	1.040	1.061	1.083	1.106	1.130	1.155	1.181			
30	1.016	1.032	1.049	1.067	1.085	1.104	1.123	1.143			
50	1.012	1.025	1.038	1.051	1.064	1.078	1.093	1.107			
100	1.009	1.017	1.026	1.035	1.045	1.054	1.064	1.074			
200	1.006	1.012	1.019	1.025	1.031	1.038	1.044	1.051			

	Log-normal 5%ile Strength (Tail fit, Tight CoV)										
CL 80%		CoV (%)									
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%			
5	1.034	1.071	1.111	1.153	1.199	1.249	1.303	1.362			
10	1.024	1.049	1.076	1.104	1.133	1.164	1.197	1.232			
20	1.017	1.034	1.052	1.071	1.091	1.111	1.132	1.153			
30	1.014	1.028	1.042	1.057	1.073	1.089	1.105	1.122			
50	1.011	1.021	1.033	1.044	1.055	1.067	1.079	1.092			
100	1.007	1.015	1.023	1.031	1.039	1.047	1.055	1.063			
200	1.005	1.011	1.016	1.021	1.027	1.033	1.038	1.044			

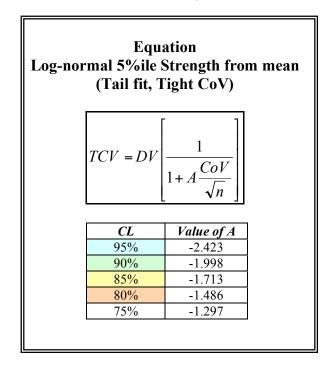
	Log-normal 5%ile Strength (Tail fit, Tight CoV)										
CL 75%		CoV (%)									
No. Sample	5%	5% 10% 15% 20% 25% 30% 35% 40%									
5	1.030	1.062	1.095	1.131	1.170	1.211	1.255	1.302			
10	1.021	1.043	1.066	1.089	1.114	1.140	1.168	1.196			
20	1.015	1.030	1.045	1.062	1.078	1.095	1.113	1.131			
30	1.012	1.024	1.037	1.050	1.063	1.076	1.090	1.105			
50	1.009	1.019	1.028	1.038	1.048	1.058	1.069	1.079			
100	1.007	1.013	1.020	1.027	1.033	1.040	1.048	1.055			
200	1.005	1.009	1.014	1.019	1.023	1.028	1.033	1.038			

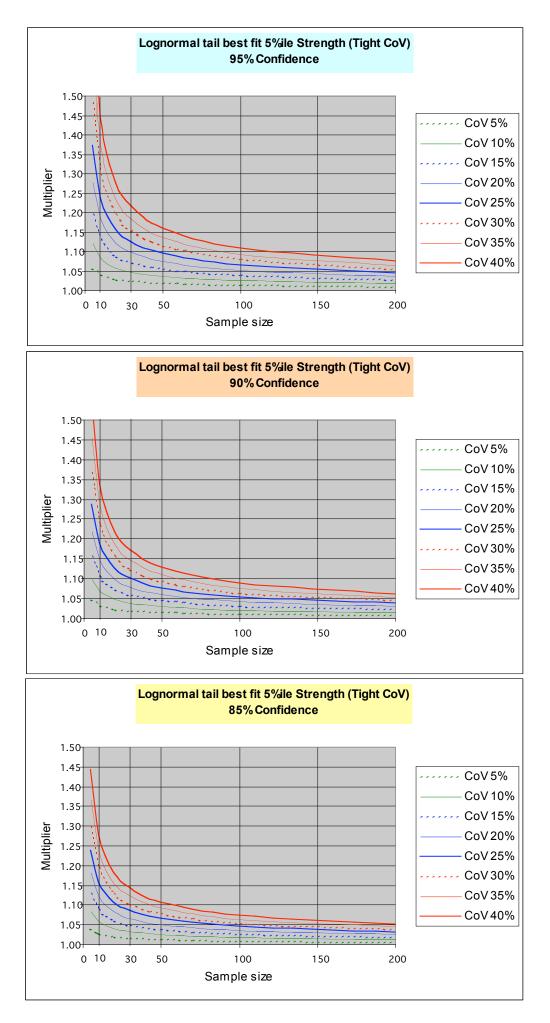
The estimation of 5% ile strength using a log-normal distribution through the tail of the data makes use of the natural logarithms of the raw strength data (ln(strength)). Extra columns need to be introduced to an analysis spreadsheet, to accommodate ln(strength) and probability.

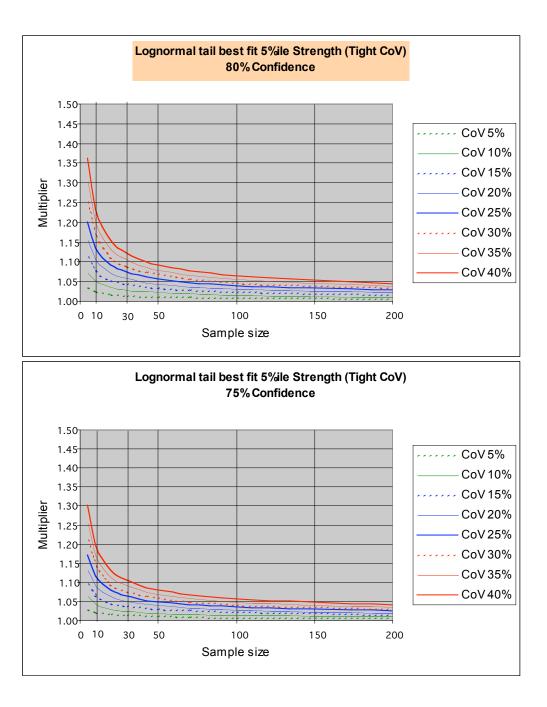
- The data is ranked and probabilities (*pr*) determined in accordance with Appendix C.2.5.
- The standard deviation of the log-normal data is found from (*eqn C.12*) using the estimated CoV of the product.
- The mean of the log-normal distribution is found from the ln(strength) and *pr* data using (*eqn C.14*)
- The 5% ile strength can be found using (eqn C.11).
- The LN, EXP and AVERAGE functions in msExcel are used.

The Test Comparison Value used to accept or reject this estimate is found using the following equation or the preceding Tables.

This equation is used when the estimated CoV of production is within 5% of the long-term CoV (ie. Between 0.95 CoV and 1.05 CoV).







B.7 Multipliers for Log-normal 5%ile Strength from mean (Tail fit, Loose CoV)

$$TCV = M * DV$$

	Log-normal 5%ile Strength (Tail fit, Loose CoV)										
CL 95%		CoV (%)									
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%			
5	1.071	1.152	1.247	1.359	1.492	1.655	1.859	2.118			
10	1.049	1.103	1.163	1.229	1.304	1.389	1.485	1.596			
20	1.034	1.071	1.110	1.152	1.198	1.247	1.300	1.359			
30	1.028	1.057	1.088	1.121	1.156	1.193	1.232	1.275			
50	1.021	1.044	1.067	1.091	1.116	1.143	1.171	1.200			
100	1.015	1.030	1.046	1.063	1.080	1.097	1.115	1.134			
200	1.011	1.021	1.032	1.044	1.055	1.067	1.079	1.091			

	Log-normal 5%ile Strength (Tail fit, Loose CoV)									
CL 90%		CoV (%)								
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%		
5	1.060	1.127	1.204	1.292	1.393	1.512	1.653	1.823		
10	1.042	1.087	1.136	1.190	1.249	1.315	1.388	1.469		
20	1.029	1.060	1.092	1.127	1.164	1.204	1.246	1.292		
30	1.024	1.048	1.074	1.102	1.130	1.160	1.192	1.226		
50	1.018	1.037	1.057	1.077	1.098	1.120	1.143	1.167		
100	1.013	1.026	1.039	1.053	1.067	1.082	1.097	1.112		
200	1.009	1.018	1.028	1.037	1.047	1.057	1.067	1.077		

	Log-normal 5%ile Strength (Tail fit, Loose CoV)									
CL 85%		CoV (%)								
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%		
5	1.053	1.111	1.177	1.250	1.334	1.429	1.539	1.668		
10	1.037	1.076	1.119	1.165	1.215	1.270	1.329	1.395		
20	1.026	1.053	1.081	1.111	1.143	1.177	1.212	1.250		
30	1.021	1.043	1.065	1.089	1.114	1.140	1.167	1.195		
50	1.016	1.033	1.050	1.068	1.086	1.105	1.125	1.145		
100	1.011	1.023	1.035	1.047	1.059	1.072	1.085	1.098		
200	1.008	1.016	1.024	1.033	1.041	1.050	1.059	1.068		

	Log-normal 5%ile Strength (Tail fit, Loose CoV)										
CL 80%		CoV (%)									
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%			
5	1.047	1.099	1.156	1.219	1.290	1.370	1.459	1.562			
10	1.033	1.068	1.105	1.146	1.189	1.236	1.286	1.341			
20	1.023	1.047	1.072	1.099	1.127	1.156	1.187	1.219			
30	1.019	1.038	1.058	1.079	1.101	1.124	1.147	1.172			
50	1.014	1.029	1.045	1.060	1.077	1.093	1.111	1.128			
100	1.010	1.021	1.031	1.042	1.053	1.064	1.076	1.087			
200	1.007	1.014	1.022	1.029	1.037	1.045	1.052	1.060			

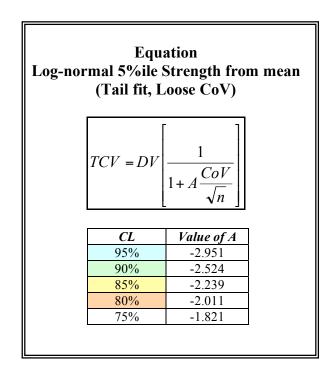
	Log-normal 5%ile Strength (Tail fit, Loose CoV)										
CL 75%		CoV (%)									
No. Sample	5%	10%	15%	20%	25%	30%	35%	40%			
5	1.042	1.089	1.139	1.195	1.256	1.323	1.399	1.483			
10	1.030	1.061	1.095	1.130	1.168	1.209	1.252	1.299			
20	1.021	1.042	1.065	1.089	1.113	1.139	1.166	1.195			
30	1.017	1.034	1.052	1.071	1.091	1.111	1.132	1.153			
50	1.013	1.026	1.040	1.054	1.069	1.084	1.099	1.115			
100	1.009	1.019	1.028	1.038	1.048	1.058	1.068	1.079			
200	1.006	1.013	1.020	1.026	1.033	1.040	1.047	1.054			

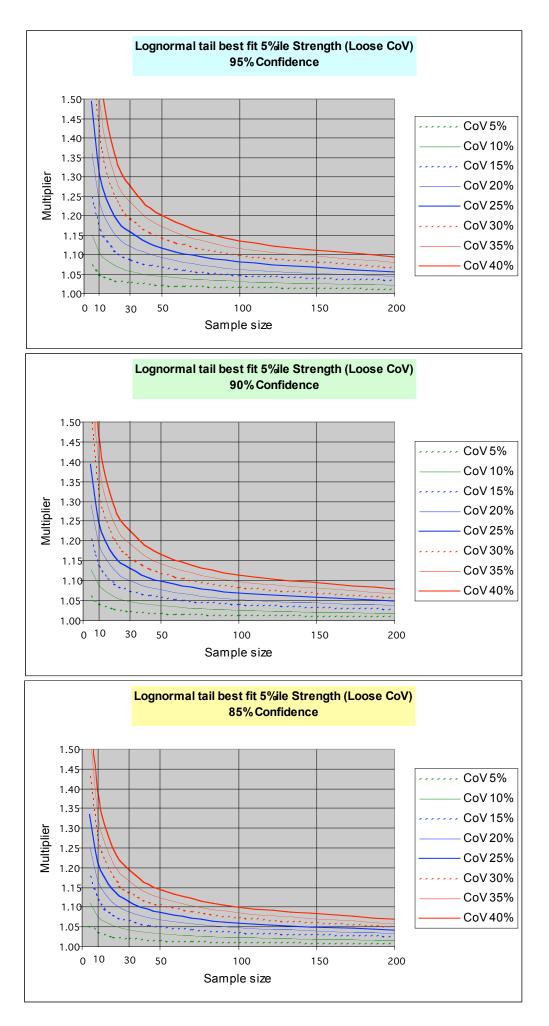
The estimation of 5% ile strength using a log-normal distribution through the tail of the data makes use of the natural logarithms of the raw strength data (ln(strength)). Extra columns need to be introduced to an analysis spreadsheet, to accommodate ln(strength) and probability.

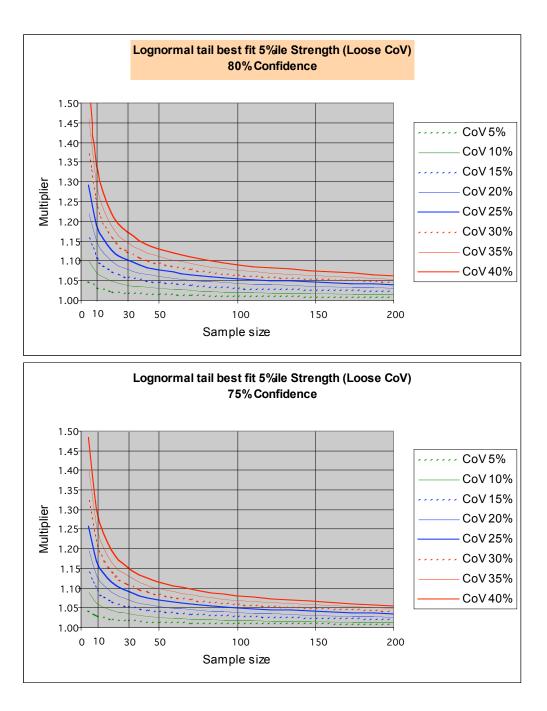
- The data is ranked and probabilities (*pr*) determined in accordance with Appendix C.2.5.
- The standard deviation of the log-normal data is found from (*eqn C.12*) using the estimated CoV of the product.
- The mean of the log-normal distribution is found from the ln(strength) and *pr* data using (*eqn C.14*)
- The 5% ile strength can be found using (*eqn C.11*).
- The LN, EXP and AVERAGE functions in msExcel are used.

The Test Comparison Value used to accept or reject this estimate is found using the following equation or the preceding Tables.

This equation is used when the estimated CoV of production is within 10% of the long-term CoV (ie. Between 0.90 CoV and 1.10 CoV).







B.8 Multipliers for Two Parameter Weibull (Tail fit) 5%ile Strength

2 Parameter Weibull (Tail fit) 5%ile Strength									
CL 95%		CoV (%)							
Sample size	5%	10%	15%	20%	25%	30%	35%	40%	
30	1.061	1.130	1.208	1.298	1.403	1.526	1.673	1.851	
50	1.047	1.098	1.154	1.217	1.286	1.364	1.453	1.553	
100	1.032	1.067	1.104	1.144	1.187	1.233	1.283	1.337	
200	1.023	1.047	1.072	1.098	1.125	1.154	1.185	1.217	

2 Parameter Weibull (Tail fit) 5%ile Strength									
CL 90%		CoV (%)							
Sample size	5%	10%	15%	20%	25%	30%	35%	40%	
30	1.049	1.102	1.162	1.228	1.302	1.386	1.481	1.590	
50	1.037	1.077	1.121	1.168	1.219	1.275	1.336	1.404	
100	1.026	1.054	1.083	1.113	1.146	1.180	1.216	1.255	
200	1.018	1.037	1.057	1.077	1.099	1.121	1.144	1.168	

	2 Parameter Weibull (Tail fit) 5%ile Strength							
CL 85%		CoV (%)						
Sample size	5%	10%	15%	20%	25%	30%	35%	40%
30	1.041	1.085	1.133	1.186	1.243	1.307	1.377	1.456
50	1.031	1.065	1.100	1.138	1.179	1.222	1.269	1.320
100	1.022	1.045	1.069	1.094	1.120	1.148	1.176	1.207
200	1.015	1.031	1.048	1.065	1.082	1.100	1.119	1.138

	2 Parameter Weibull (Tail fit) 5%ile Strength							
CL 80%		CoV (%)						
Sample size	5%	10%	15%	20%	25%	30%	35%	40%
30	1.034	1.071	1.111	1.153	1.200	1.249	1.304	1.363
50	1.026	1.054	1.084	1.115	1.148	1.183	1.220	1.260
100	1.019	1.038	1.058	1.079	1.100	1.123	1.146	1.171
200	1.013	1.026	1.040	1.054	1.069	1.084	1.099	1.115

	2 Parameter Weibull (Tail fit) 5%ile Strength							
CL 75%		CoV (%)						
Sample size	5%	10%	15%	20%	25%	30%	35%	40%
30	1.029	1.060	1.092	1.127	1.164	1.203	1.245	1.291
50	1.022	1.046	1.070	1.096	1.122	1.150	1.180	1.211
100	1.016	1.032	1.048	1.066	1.084	1.102	1.121	1.141
200	1.011	1.022	1.034	1.046	1.058	1.070	1.083	1.096

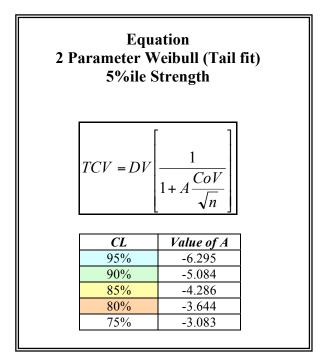
The estimation of 5% ile strength using a 2 parameter Weibull distribution through the tail of the data makes use of the natural logarithms of the raw strength data (ln(strength)). Extra columns need to be introduced to an analysis spreadsheet, to accommodate ln(strength) and probability.

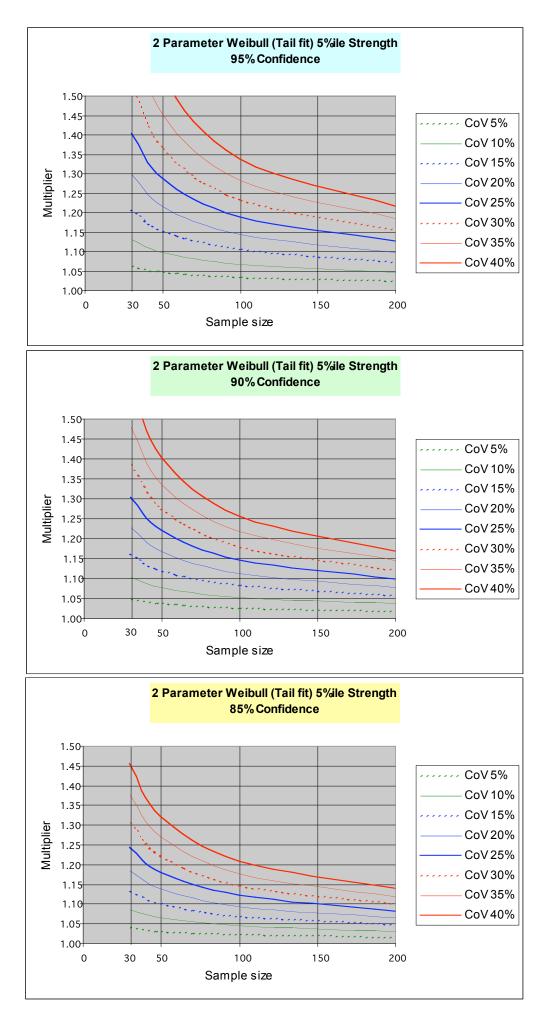
- The data is ranked and probabilities (*pr*) determined in accordance with Appendix C.2.6.
- The log(strength) is plotted against a function of *pr* as indicated in Appendix C.2.6.
- The shape and scale parameters can be found from the slope and intercept of the plot.
- The 5% ile strength can be found using (*eqn C.17*).
- The LN, EXP, SLOPE and INTERCEPT functions in msExcel are used.

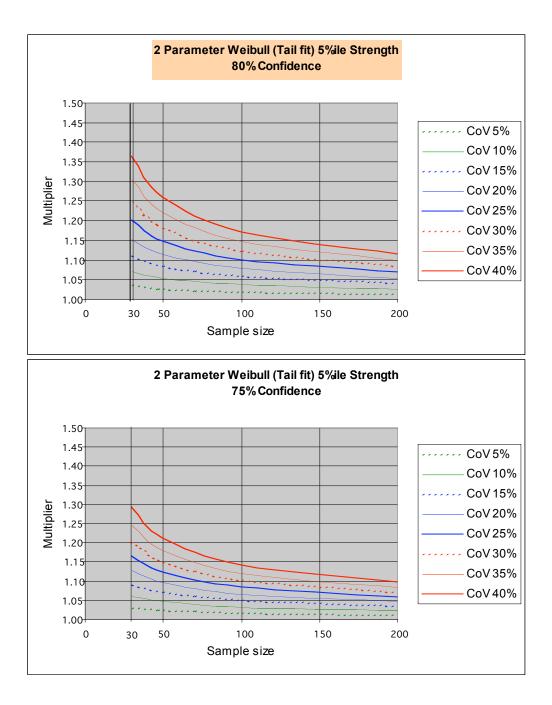
The Test Comparison Value used to accept or reject this estimate is found using the following equation or the preceding Tables.

The relationships were established using a proof stress equivalent to the 15% ile of the whole data.

Note: This method is valid for sample sizes greater than 30.







B.9 Multipliers for Two Parameter Weibull (Tail fit) ISO 13910 5%ile Strength

TCV =	М	*	DV
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	2 Parameter Weibull (Tail fit) ISO 13910 5%ile Strength							
CL 95%		CoV (%)						
Sample size	5%	10%	15%	20%	25%	30%	35%	40%
30	1.029	1.060	1.093	1.128	1.165	1.205	1.248	1.293
50	1.022	1.046	1.071	1.096	1.123	1.152	1.182	1.213
100	1.016	1.032	1.049	1.066	1.084	1.103	1.122	1.142
200	1.011	1.022	1.034	1.046	1.058	1.071	1.083	1.096

	2 Parameter Weibull (Tail fit) ISO 13910 5%ile Strength							
CL 90%		CoV (%)						
Sample size	5%	10%	15%	20%	25%	30%	35%	40%
30	1.023	1.046	1.071	1.097	1.124	1.153	1.183	1.215
50	1.017	1.035	1.054	1.073	1.094	1.114	1.136	1.158
100	1.012	1.025	1.038	1.051	1.064	1.078	1.092	1.107
200	1.009	1.017	1.026	1.035	1.045	1.054	1.064	1.073

	2 Parameter Weibull (Tail fit) ISO 13910 5%ile Strength							
CL 85%		CoV (%)						
Sample size	5%	10%	15%	20%	25%	30%	35%	40%
30	1.018	1.037	1.056	1.077	1.098	1.120	1.142	1.166
50	1.014	1.028	1.043	1.058	1.074	1.090	1.107	1.124
100	1.010	1.020	1.030	1.041	1.051	1.062	1.073	1.085
200	1.007	1.014	1.021	1.028	1.036	1.043	1.051	1.058

	2 Parameter Weibull (Tail fit) ISO 13910 5%ile Strength							
CL 80%		CoV (%)						
Sample size	5%	10%	15%	20%	25%	30%	35%	40%
30	1.015	1.030	1.045	1.061	1.078	1.095	1.112	1.130
50	1.011	1.023	1.035	1.047	1.059	1.072	1.085	1.098
100	1.008	1.016	1.024	1.033	1.041	1.050	1.058	1.067
200	1.006	1.011	1.017	1.023	1.029	1.035	1.041	1.047

	2 Parameter Weibull (Tail fit) ISO 13910 5%ile Strength							
CL 75%		CoV (%)						
Sample size	5%	10%	15%	20%	25%	30%	35%	40%
30	1.012	1.024	1.036	1.048	1.061	1.074	1.088	1.101
50	1.009	1.018	1.027	1.037	1.047	1.056	1.067	1.077
100	1.006	1.013	1.019	1.026	1.033	1.039	1.046	1.053
200	1.004	1.009	1.014	1.018	1.023	1.027	1.032	1.037

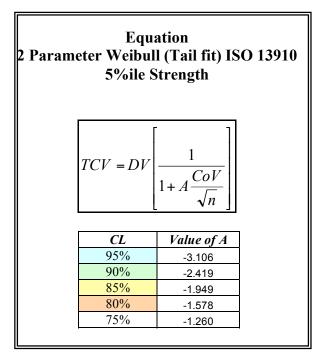
The estimation of 5% ile strength using a 2 parameter Weibull distribution through the tail of the data makes use of the natural logarithms of the raw strength data (ln(strength)). Extra columns need to be introduced to an analysis spreadsheet, to accommodate ln(strength) and probability.

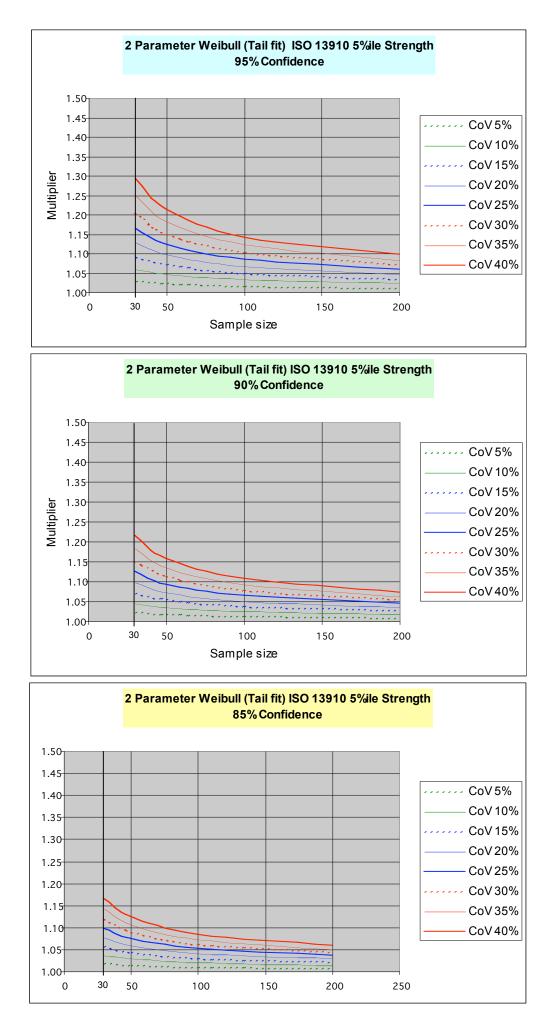
- The data is ranked and probabilities (*pr*) determined in accordance with Appendix C.2.6. The tail of the distribution used in the analysis has the greater of 15 points or 15% of the sample. The lowest two points are discarded in the remainder of the analysis.
- The log(strength) is plotted against a function of *pr* as indicated in Appendix C.2.6.
- The shape and scale parameters can be found from the slope and intercept of the plot.
- The 5% ile strength can be found using (*eqn C.17*).
- The LN, EXP, SLOPE and INTERCEPT functions in msExcel are used.

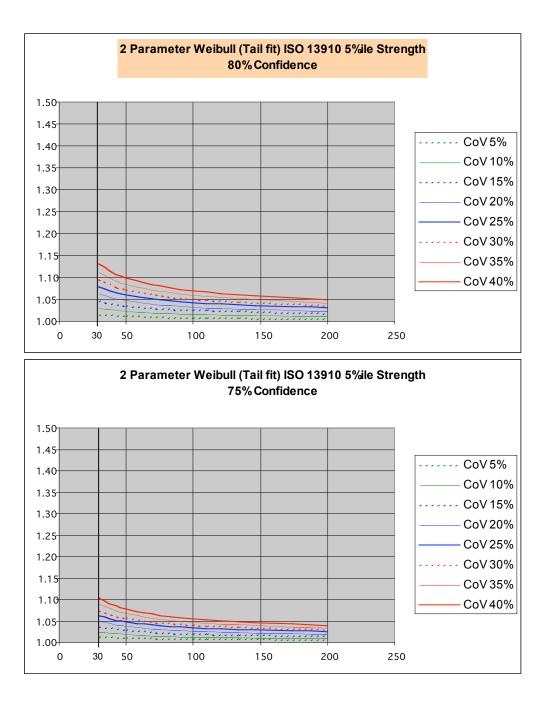
The Test Comparison Value used to accept or reject this estimate is found using the following equation or the preceding Tables.

The relationships were established using a proof stress equivalent to the 15% ile of the whole data.

Note: This method is valid for sample sizes greater than 30.







Appendix C – Technical Support

C.1 Confidence Limits

C.1.1 Confidence in mean of Normal distribution

For normal distributions, the expression for a confidence limit on an estimation of the mean is given in $(eqn \ C.1)^{1,2}$.

$$CL(pr) = \mu + z(pr)\frac{\sigma}{\sqrt{n}}$$
 (eqn C.1)

with

CL(pr)	Confidence limit of estimate of mean
(<i>pr</i>)	Confidence level (eg 95%)
μ	Mean of sample (size <i>n</i>)
z(pr)	Normal distribution (a function of probability
	level)
σ	Standard deviation
N	Sample size (number of specimens in a single
	sample)
CoV	Coefficient of Variation = σ/μ

Then using
$$CoV = \frac{\sigma}{\mu}$$
 then $\sigma = \mu CoV$

And
$$CL(pr) = \mu + z(pr) \frac{\mu CoV}{\sqrt{n}}$$

Hence $CL(pr) = \mu \left(1 + z(pr) \frac{CoV}{\sqrt{n}}\right)$ (eqn C.2)

C.1.2 Confidence in any statistical distribution

(eqn C.2) can be generalised to other distributions (other than normal distribution) and other estimators (other than the mean – more importantly the 5% ile)¹⁰.

$$CL(pr) = X\left(1 + A\frac{CoV}{\sqrt{n}}\right)$$
 (eqn C.3)

with

A	Constant for each Confidence level
X	Estimated property for the population

This is the origin of $(eqn \ C.24)$ in this report. The values of A appropriate to each estimator¹¹ and method of analysis were determined using Monte Carlo simulations.

C.1.3 Design Values and Test Comparison Values

In AS/NZS 4063¹⁰ and most other similar standards in other countries, a Design Value is found from a given confidence limit on a test value. Here we can make the substitution:

• DV = CL(pr) – the Design Value is the confidence limit on the test value X.

This then gives (eqn C.3) as

$$DV = X \left(1 + A \frac{CoV}{\sqrt{n}} \right)$$

and $X_1 = DV \left[\frac{1}{\left(1 + A \frac{CoV}{\sqrt{n}} \right)} \right]$

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Here X_1 is the value of the test result that has the given confidence of being greater than the Design Value (DV). This is defined as the Test Comparison Value, giving

$$TCV = DV \left[\frac{1}{\left(1 + A \frac{CoV}{\sqrt{n}} \right)} \right]$$
 (eqn C.4)

with

TCV	Test Comparison Value (test value with given
	level of confidence that it meets the Design Value)
DV	Design Value
A	Constant for each Confidence level and each
	analysis method
n	Sample size (number of specimens in a single
	sample)
CoV	Coefficient of Variation = σ/μ

This is the origin of (eqn 5.4) and the multiplier in the square brackets in (eqn C.4) is the value *M* which is presented in all of the tables in Appendices A and B.

C.2 Methods of analysis

C.2.1 Non-parametric estimates

For non-parametric estimates from data, no attempt is made to fit a distribution through the data, it is simply treated as a list of numbers. It is outlined in Section 4.2.1

Non-parametric studies can make use of standard functions in msExcel:

- AVERAGE(multiple cell reference) returns the result of (eqn C.5)
- STDEV(multiple cell reference) returns the result of (*eqn C.7*)
- PERCENTILE(multiple cell reference, 0.05) gives an answer close to the value of the 5% ile calculated using the method outlined here.

The following nomenclature applies:

X _i	A test value. It is a typical value. The first value would be X_1 , the second X_2 , and a general one (the i^{th} one) X_i .
$\sum X_i$	The sum of all test values
\overline{X}	The arithmetic average of all test values
$X_{0.05}$	5 th percentile of the test values
S_X	The standard deviation of all test values
n	Sample size (number of specimens in a single sample)
CoV	Coefficient of Variation – see (eqn C.6)

Estimating the mean

The mean is the arithmetic average of all of the test data.

$$\overline{X} = \frac{\sum X_i}{n}$$
 (eqn C.5)

Estimating the Coefficient of Variation (CoV)

The Coefficient of Variation (CoV) is the ratio of standard deviation to the average.

$$CoV = \frac{s_x}{\overline{X}}$$
(eqn C.6)
$$s_x = \sqrt{\frac{\sum \left(X_i - \overline{X}\right)^n}{n}}$$
(eqn C.7)

Estimating the 5% ile ($X_{0.05}$)

with

The non-parametric estimate of the 5% ile is simply an interpolation between the ranked data to give the value at the 5% ile level. Section 4.2.1 shows an example.

- All of the data is ranked (ordered from smallest to largest)
- The probability of each data point can be found from its ranking.
 - The first lowest ranked value, has $pr = \frac{1 0.5}{n}$

• The second value, has
$$pr = \frac{2 - 0.5}{n}$$

• The *i*th value, has $pr = \frac{i - 0.5}{n}$

- The data point with the *pr* just greater than 0.05 is found $(X_{0.05+})$, and the one with the *pr* just less than 0.05 is found $(X_{0.05-})$.
- The value of $X_{0.05}$ is found by linear interpolation between $X_{0.05+}$ and $X_{0.05-}$.

C.2.2 Log-normal estimates

As outlined in Section 4.2.2, log-normal estimates can be found by fitting a lognormal distribution through the data, and finding critical parameters from the distribution that is fitted.

Fitting the distribution

The natural logarithm is found for all of the data.

$$v_i = \ln(X_i)$$

(*eqn C*.8)

All of the subsequent operations are performed on the transformed data, using two parameters derived from the set of ln(X).

Mean of logs

$$m = \frac{\sum v_i}{n}$$

s $s_v = \sqrt{\frac{\sum (v_i - m)^2}{n}}$

Standard deviation of logs	$s_v = 1$

X _i	A test value. It is a typical value. The first value
	would be X_1 , the second X_2 , and a general one (the i^{th} one) X_i .
v_i	The natural log of the test values.
$\sum v_i$	The sum of all <i>v</i> values
т	The arithmetic average of all <i>v</i> values
$X_{0.05}$	5 th percentile of the log-normal distribution
	through the test values – see (<i>eqn C.11</i>)
S_{V}	The standard deviation of all <i>v</i> values
n	Sample size (number of specimens in a single
	sample)
CoV	Coefficient of Variation of the distribution through
	the data – see $(eqn \ C.10)$
exp()	The exponentiation function $\exp(s) = e^s$

Estimating the mean

The mean of the data is estimated from the distribution itself rather than the test values. It uses the values of m and s_v defined above:

$$\overline{X} = \exp\left(m + \frac{s_v^2}{2}\right) \tag{eqn C.9}$$

Estimating the Coefficient of Variation (CoV)

Again, the Coefficient of Variation (CoV) of the data can be estimated from the lognormal distribution fitted through the data rather than from the data itself. It uses the values of m and s_v defined above:

$$CoV = \sqrt{\exp(\mathfrak{s}_{v}^{2})} - 1 \qquad (eqn \ C.10)$$

Estimating the 5% ile $(X_{0.05})$

The 5% ile of the data can be estimated from the log-normal distribution fitted through the data rather than from the data itself. It uses the values of *m* and s_v defined above:

$$X_{0.05} = \exp(m - 1.645s_{\nu})$$
 (eqn C.11)

C.2.3 Log-normal distribution estimates from tail data only

The fitting method outlined in Section C.2.2 required all of the test data, so could not be fitted to data from a proof testing program. However, the log-normal distribution can be fitted through just the tail data as shown in this section. Once the values of m and s_v have been found, then (eqn C.9), (eqn C.10) and (eqn C.11) can be used to find the estimates of mean, CoV and 5% ile of the data respectively.

Fitting the distribution

- The data is ranked from lowest to highest for which test information is available.
- The probability is assigned to each data point using the same relationships outlined in Section C.2.1. Here *n* is the total number of pieces where a test was started, not the number for which data was available. (eg if 30 pieces were loaded, but only 6 broke, then there is data for 6, so only 6 pieces can be ranked, but n = 30)

• The first – lowest ranked – value, has
$$pr = \frac{1 - 0.5}{n}$$

• The second value, has
$$pr = \frac{2 - 0.5}{n}$$

• The *i*th value, has
$$pr_i = \frac{i - 0.5}{n}$$

- For each data point, $v_i = \ln(X_i)$ is plotted against $z(pr_i) =$ normal distribution points for each probability. The line of best fit is fitted through the plot of the data.
 - The value of m is given by the intercept on the v axis.
 - The value of s_v is given by the slope of the line of best fit.

C.2.4 Log-normal distribution from the mean of the data with specified CoV

In Section 4.2.3, it was noted that in fitting a distribution to the data, the character of the data may reflect sampling errors rather than the distribution of production properties from which the sample was drawn. It is possible to reflect the character of the population better by using the known CoV of the production property rather than the one suggested by the data from just a single sample.

In common with all log-normal distributions, the data is transformed to the natural logarithms using (eqn C.8).

$$v_i = \ln(X_i)$$
 (copy eqn C.8)

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(*eqn C*.10) related *CoV* and the s_v value for the log-normal distribution. By inverting this expression, the known population value of CoV can give the required s_v directly.

$$CoV = \sqrt{\exp(s_{v}^{2}) - 1} \qquad (copy \ eqn \ C.10)$$

$$CoV^{2} = \exp(s_{v}^{2}) - 1$$

$$\exp(s_{v}^{2}) = CoV^{2} + 1$$

$$s_{v}^{2} = \ln(CoV^{2} + 1)$$

$$s_{v} = \sqrt{\ln(CoV^{2} + 1)} \qquad (eqn \ C.12)$$

X_i	A test value. It is a typical value. The first value
	would be X_1 , the second X_2 , and a general one (the
	i^{th} one) X_i .
v_i	The natural log of the test values.
$\sum v_i$	The sum of all <i>v</i> values
т	The arithmetic average of all <i>v</i> values
$X_{0.05}$	5 th percentile of the log-normal distribution
	through the test values – see (eqn C.11)
S_V	The standard deviation of all <i>v</i> values
n	Sample size (number of specimens in a single
	sample)
CoV	Coefficient of Variation of the distribution through
	the data – see $(eqn \ C.10)$
z(pr)	Points on the normal distribution – function of
· · /	probability
exp()	The exponentiation function eg. $exp(s) = e^s$
UAPU	The exponentiation function eg. $exp(s) = e$

 s_v is calculated using only the *CoV* that is expected for the grade of timber that has been sampled. The value of *m* can be found from the test data and the value of s_v found from (*eqn C.12*).

- The data is ranked from lowest to highest.
- The probability is assigned to each data point using the same relationships outlined in Section C.2.1.

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• The first – lowest ranked – value, has
$$pr = \frac{1-0.5}{n}$$

• The second value, has $pr = \frac{2-0.5}{n}$
• The *i*th value, has $pr_i = \frac{i-0.5}{n}$

• *m* is evaluated using (*eqn C.13*)

$$m = \frac{\sum v_i}{n} - s_v \frac{\sum z(pr_i)}{n}$$
 (eqn C.13)

with $z(pr_i)$ the normal distribution points for probability of each point (pr_i)

Once the values of *m* and s_v have been found, then (*eqn C.9*), (*eqn C.10*) and (*eqn C.11*) can be used to find the estimates of mean, *CoV* and 5% ile of the data respectively.

In using this method of analysis, a single value of CoV is used, but because, there will be a small variation in CoV of the production from time to time, the actual value of the CoV of production may differ a little from the value used in the analysis. In modelling the analysis method, allowance is made for the variation in CoV (see Section C.3.3).

C.2.5 Log-normal distribution from the mean of the data with specified CoV using tail fit only

With this method of evaluation of the critical properties of the data, a log-normal distribution is fitted through the data points available, but the distribution is forced to have the CoV known for the grade produced.

The method is similar to the one presented in Section C.2.4. In common with all lognormal distributions, the data is transformed to the natural logarithms using (eqnC.8). $v_i = \ln(X_i)$ (copy eqn C.8)

 s_v is calculated using only the *CoV* that is expected for the grade of timber that has been sampled using (*eqn C.12*).

$$s_v = \sqrt{\ln(CoV^2 + 1)} \qquad (\text{copy } eqn \ C.12)$$

The value of *m* can be found from the test data and the value of s_v found from (eqn C.12).

- The data is ranked from lowest to highest for which test information is available. Here *n* is the total number of pieces where a test was started, not the number for which data was available. (eg if 30 pieces were loaded, but only 6 broke, then there is data for 6, so only 6 pieces can be ranked, but n = 30)
- The probability is assigned to each data point using the same relationships outlined in Section C.2.1.
 - The first lowest ranked value, has $pr = \frac{1 0.5}{n}$

• The second value, has
$$pr = \frac{2 - 0.5}{n}$$

• The *i*th value, has
$$pr_i = \frac{i - 0.5}{n}$$

• *m* is evaluated using (*eqn* C.14) but in this case, the divisor is n_b – the number of pieces that broke and for which strength information is available.

$$m = \frac{\sum v_i}{n_b} - s_v \frac{\sum z(pr_i)}{n_b}$$
 (eqn C.14)

with $z(pr_i)$ the normal distribution points for probability of each point (pr_i)

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Once the values of *m* and s_v have been found, then (*eqn C.9*), (*eqn C.10*) and (*eqn C.11*) can be used to find the estimates of mean, *CoV* and 5% ile of the data respectively.

In using this method of analysis, a single value of CoV is used, but because, there will be a small variation in CoV of the production from time to time, the actual value of the CoV of production may differ a little from the value used in the analysis. In modelling the analysis method, allowance is made for the variation in CoV (see Section C.3.3).

C.2.6 Two parameter Weibull estimates

Tail fitting a 2 parameter Weibull (2pW) distribution is used in US standards and has been adopted in a number of mills for monitoring programs. Under the US standard, the tail is defined as the larger of (the lowest 15% or the lowest 15 pieces).

The tail fit 2 parameter Weibull distribution appears to have a good match for the tail data, but gives a CoV that is significantly lower than the CoV of the whole distribution. For sample sizes of less than 100, the definition of the tail gives 15 pieces as the limiting criteria, and that will represent something more than 15% of the data. This will change the ratio of the 2pW CoV to the whole of data CoV. The errors in this slight discrepancy were not significant for sample sizes of 30 and of 50, but they were for sample sizes less than 30. For this reason, the study restricted the use of tail-fit data for the 2pW distribution to samples of 30 or more pieces.

The 2 parameter Weibull distribution is characterised by two parameters that have a variety of symbols. In this report, the symbols *a* and *b* will be used to define the distribution.

- The shape parameter = 1/a
- The scale parameter = $\exp(b)$

Once these parameters have been evaluated, then the mean, CoV and 5% ile can be estimated from the distribution using a and b^{16} .

- The data is ranked from lowest to highest for which test information is available. Here *n* is the total number of pieces where a test was started, not the number for which data was available. (eg if 150 pieces were loaded, but only 23 are used as the tail, then only the lowest 23 pieces are used, but n = 150)
- The probability is assigned to each data point using the same relationships outlined in Section C.2.1.

• The first – lowest ranked – value, has
$$pr = \frac{1 - 0.5}{n}$$

• The second value, has
$$pr = \frac{2 - 0.5}{n}$$

• The *i*th value, has
$$pr_i = \frac{i - 0.5}{n}$$

- For each point the natural logarithm of the strength is used (eqn C.8) and this is plotted against $\ln(-\ln(1-pr_i))$.
 - The slope of this plot gives *a*.
 - The intercept on the v axis gives b.

r	
v_i	The natural log of the test values.
a	A parameter of 2 parameter Weibull distribution
	shape factor = $1/a$
b	A parameter of 2 parameter Weibull distribution
	scale factor = $\exp(b)$
$X_{0.05}$	5 th percentile of the log-normal distribution
	through the test values $- \sec(eqn \ C.17)$
n	Sample size (number of specimens in a single
	sample)
CoV	Coefficient of Variation of the distribution through
	the data – see (eqn $C.16$)
Π()	Gamma function
exp()	The exponentiation function eg. $exp(s) = e^s$

The calculation of the mean and standard deviation of the distribution both make use of the gamma function $\boldsymbol{\Gamma}$

Mean $\overline{X} = \exp(b)\Gamma(a+1)$ (eqn C.15) This can be found in msExcel by EXP(b)*EXP(GAMMALN(a+1))

CoV
$$CoV = \frac{\sqrt{\Gamma(2a+1) - \Gamma(a+1)^2}}{\Gamma(a+1)}$$
 (eqn C.16)
In msExcel (EXP(GAMMALN(2a+1))-EXP(GAMMALN((a+1)^2))^0.5/
(EXP(GAMMALN(a+1)))

5%ile
$$X_{0.05} = \exp(b) \times (-\ln(1 - pr_i))^a$$
 (eqn C.17)

ISO 13910 includes an example of a 2 parameter Weibull fit to tail data for the estimation of 5% ile strength. It is very similar to the above example, except that it discards the lowest two values.

- The data is ranked from lowest to highest for which test information is available. Here *n* is the total number of pieces where a test was started, not the number for which data was available. (eg if 150 pieces were loaded, but only 23 are used as the tail, then only the lowest 23 pieces are used, but *n* = 150)
- The probability is assigned to each data point using the same relationships outlined in Section C.2.1.
 - The first lowest ranked value, has $pr = \frac{1 0.5}{n}$

• The second value, has
$$pr = \frac{2 - 0.5}{n}$$

• The *i*th value, has
$$pr_i = \frac{i - 0.5}{n}$$

- The data for the lowest two points are then ignored so the data starts with the point $pr = \frac{3-0.5}{n} = \frac{2.5}{n}$
- For each point the natural logarithm of the strength is used (eqn C.8) and this is plotted against $\ln(-\ln(1-pr_i))$.
 - The slope of this plot gives *a*.
 - The intercept on the v axis gives b.

The analysis proceeds as indicated above, but the removal of the lowest two points seems to give this method more stability.

C.3 Monte Carlo Simulations

The process of Monte Carlo Simulation is one of repeating a number of analyses of data drawn at random from distributions of "raw data". (Its name is derived from the code name of similar studies performed in World War II.)

C.3.1 Monte Carlo simulations for this study

In the context of this study, whole production runs of timber were modelled using a statistical distribution. Section 4.2 discusses the process of fitting distributions to data sets of MoE, bending strength and tension strength. Monte Carlo simulations were performed using @RISK software³ – an msExcel add in by Palisade Corp.

- The distributions with appropriate properties (*CoV*, 5%ile for strength and mean for MoE) were defined in a spreadsheet. As the distribution defined all of the timber in the production run, the characteristics of the raw data were the critical points of the input distribution.
 - The average of the production run was the average of the input distribution,
 - The CoV of the production run was the CoV of the input distribution,
 - The 5%ile was the 5%ile of the input distribution.
- For each simulation run, a number of pieces per sample was defined (eg 5, 10, 20, ...) and values were selected at random from the distributions of the production run for each piece in the sample.
- Each sample was operated on to estimate the critical points for the production (mean, CoV, 5%ile). The methods used to estimate these were described in Section C.2. For each simulation run, the properties were estimated for all of the analysis methods from the same sample, giving a direct comparison between all of the analysis methods.
- Simulation runs were repeated.
 - For strength, 2500 simulation runs were used for each sample size and CoV combination.
 - For MoE, 1000 simulation runs were used for each sample size and CoV combination.
- Each simulation gave a distribution for each estimate, so the results could be interpreted directly as confidence limits.
 - Table C.1 gives an example of the output for a simulation of MoE.
 - Table C.2 gives an example of the output for a simulation of strength.

Production Data	4
CoV of population	0.15
no of specimens	30
pop 5%ile	8.124573
Mean	10.5
std deviation	1.575

	average	Average	CoV full		5%ile full	5%ile full
	nonparm	In	nonparm	CoV full In	nonparm	In
minimum	9.566862	9.57172	0.100254	9.81E-02	6.87E+00	7.035088
maximum	11.42296	11.41863	0.231226	0.212373	9.808073	9.153103
mean	10.50024	10.5041	0.14845	0.148643	8.311314	8.150551
st deviation	0.295478	2.96E-01	1.99E-02	1.92E-02	4.12E-01	0.347304
errors	0	0	0	0	0	0
5%	10.00369	10.01177	0.117385	0.118519	7.602729	7.559521
10%	10.12494	10.1305	0.124278	0.124936	7.781636	7.688518
15%	10.19867	10.20321	0.1284	0.129082	7.895227	7.797836
20%	10.24428	10.24816	0.131826	0.132309	7.977114	7.859321
25%	10.29872	10.30053	0.134784	0.135088	8.036125	7.909787
30%	10.34139	10.34396	0.137535	0.137964	8.090206	7.963676
35%	10.3784	10.38211	0.140099	0.141036	8.166368	8.023146
40%	10.41615	10.41885	0.142563	0.143446	8.216382	8.063827
45%	10.4585	10.46219	0.144761	0.145318	8.272581	8.114374
50%	10.49579	10.49927	0.146606	0.147627	8.323778	8.150947
55%	10.53401	10.53872	0.149173	0.150247	8.380261	8.192926
60%	10.57351	10.57658	0.152212	0.152448	8.42844	8.238899
65%	10.61409	10.61863	0.154719	0.154686	8.47611	8.287342
70%	10.65934	10.66178	0.157611	0.157534	8.529819	8.345545
75%	10.69678	10.70255	0.160811	0.160221	8.585226	8.399123
80%	10.7553	10.75742	0.163729	0.163868	8.66181	8.45155
85%	10.8125	10.81488	0.168346	0.168718	8.731902	8.512689
90%	10.8708	10.8732	0.174406	0.173815	8.843313	8.590223
95%	11.00238	11.00884	0.184841	0.182686	8.974718	8.702073

The data in the 5% row gives the 5% ile probability of the estimate of the mean MoE or 5% ile MoE.

Likewise the data in the 10% row gives the 10% ile probability, the data in the 15% row gives the 15% probability, etc.

The data in the row headed "mean" gives the average of all of the estimations. Where it is higher than the true average of the production (in this case, 10.5 GPa), there is a systematic over-estimation of the property.

For the strength evaluation, the table was a little different because some of the analysis methods used proof testing and for those methods, the proof stress was also an input, as it established the tail of the distribution. The proof stress is also passed to the analysis as shown in Table C.2.

Production Data	10
CoV of population	0.35
CoV of tail	0.171939
no of specimens	20
proof stress	25
pop 5%ile	20
Mean	37.06454
std deviation	12.97259
iterations	2500

Table C 2 Simulation	Output for strongth	(production CoV 35%)
Table C.2 Simulation	Output for strength	(production CoV 35%)

							5%ile tail
		5%ile full	5%ile full	5%ile tail	5%ile tail	5%ile In fr	2pW best
	count tail	nonparm	In	In best fit	In CoV	mean CoV	fit
minimum	0	12.0123	13.18462	12.23558	13.10572	15.37462	12.65374
maximum	10	31.51909	28.99482	24.79607	27.78092	26.48254	24.80392
mean	3.2292	21.71755	20.28748	19.90126	20.10625	20.05742	20.01205
st deviation	1.614221	2.82568	2.375791	2.24854	2.585688	1.521345	2.210343
errors	0	0	0	336	68	0	336
5%	1	17.21329	16.56301	16.0288	16.5432	17.65149	16.22705
10%	1	18.1235	17.28318	16.92358	17.09794	18.16251	17.07866
15%	2	18.7527	17.81447	17.4691	17.54026	18.49562	17.62907
20%	2	19.29856	18.25469	17.93437	17.90624	18.75236	18.05557
25%	2	19.82951	18.6263	18.36271	18.22482	19.01094	18.50735
30%	2	20.22642	18.9627	18.72441	18.5883	19.23816	18.85052
35%	2	20.6048	19.30079	19.04301	18.92614	19.4678	19.17792
40%	3	20.9276	19.62585	19.36721	19.19468	19.63463	19.50288
45%	3	21.307	19.90398	19.6982	19.50868	19.82478	19.81574
50%	3	21.64615	20.1936	20.00205	19.75258	19.99656	20.11393
55%	3	21.97873	20.47451	20.2742	20.01671	20.17908	20.39157
60%	4	22.3427	20.77836	20.60618	20.3201	20.3554	20.70707
65%	4	22.71919	21.14427	20.88702	20.68761	20.59343	20.99132
70%	4	23.16552	21.4905	21.19899	21.09094	20.79559	21.2913
75%	4	23.55849	21.81982	21.5602	21.5595	21.03709	21.64373
80%	5	24.09635	22.23398	21.89019	22.1047	21.29689	21.98604
85%	5	24.61761	22.6906	22.32393	22.81242	21.62568	22.40273
90%	5	25.33358	23.38523	22.81669	23.79504	22.05161	22.88081
95%	6	26.62834	24.3536	23.51835	25.08858	22.66244	23.55066

The data in the 5% row gives the 5% probability estimate of the 5% ile strength. Likewise the data in the 10% row gives the 10% probability estimate, the data in the 15% row gives the 15% probability estimate, etc.

The data in the row headed "mean" gives the average of all of the estimations. Where it is higher than the true 5% ile of the production (in this case, 20 MPa), there is a systematic over-estimation of the property.

C.3.2 Determination of Confidence limits for analysis methods

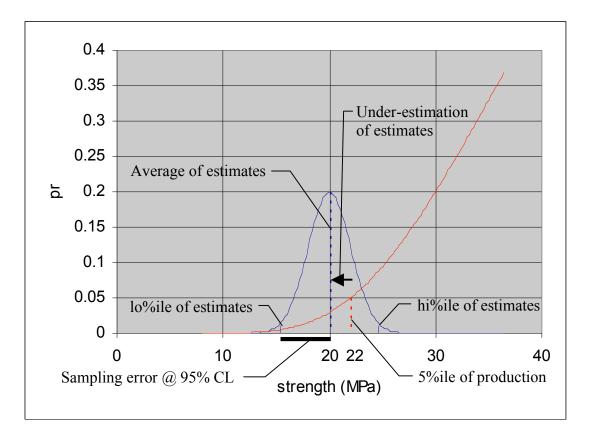
The data in Tables C.1 and C.2 represented a single combination of production CoV and sample size. In the case of the strength data shown in Table C.2, it also represented a single value of proof stress.

All of the data for a single analysis method was combined to establish values for A in (eqn 5.3) or (eqn 5.4). It is the same value of A.

Confidence level from the data

Because some analysis methods systematically over-estimated or under-estimated the production value (refer Section C.4.1), and because of slight skewness in some of the distributions of estimates, it was not straight forward to estimate the sampling error.

- Sampling errors are found for different confidence levels. (For example at the 95% confidence level, the basis for the calculations will be the 5% ile or 95% ile level of the estimations.)
- Sampling error magnitude was the difference between the percentile point of the estimation and the mean estimation of the same property.
- Over-estimation was the difference between the mean estimation and the production value of the same property.



The relationships are depicted in Figure C.1.

Figure C.1 – Representation of Estimate of production property

Figure C.1 shows the distribution of estimations of the 5% ile strength of a production. The production distribution is shown in red and the distribution of the estimate is shown in blue.

- The distribution estimate is close to the tail of the production distribution.
- In Figure C.1 there is a slight tendency to under-estimate the 5% ile. (The average estimate is 20 MPa while the true value was 22 MPa.)

• At the 95% ile level, the sampling error will be given by the 5% ile of the estimates and the 95% ile of the estimates.

For this study, the sampling error at a given confidence level (*CL*) was given by: (*upper CL value – lower CL value*)/2 (*eqn C.18*)

For this study, the over-estimation of the estimate was given by:

- (average estimation true production value) (eqn C.19)
- Where there is a systematic over- estimation this will be a positive value.
- Where there is a systematic under-estimation (as illustrated in Figure C.1), then this will be a negative value.
- This is independent of confidence level.

The lower bound on the estimate for a given Confidence Level is given by L(CL) as shown in (eqn C.20):

L(CL) = Test value - over-estimation - sampling error (CL) (eqn C.20) The over-estimation is independent of Confidence Level, and only a function of the method of analysis. However, the sampling error is both a function of the method of analysis and the Confidence Level.

In general, unless the method significantly under-estimates the property, L(CL) will be less than the true value of the parent distribution.

This can be used in (eqn 5.2) to evaluate the value of A.

$$L = X \left(1 + A \frac{CoV}{\sqrt{n}} \right) \qquad (\text{copy } eqn \ 5.2)$$

Hence

 $\frac{L}{X} = \left(1 + A\frac{CoV}{\sqrt{n}}\right) \tag{eqn C.21}$

This ratio was used to give a non-dimensional value which could be used to find A directly for a given combination of analysis method and Confidence Level.

For a given analysis method and Confidence Level, there are a number of results for different CoV and sample size combinations.

The values of $\frac{L}{X}$ could be plotted against $\frac{CoV}{\sqrt{n}}$ and the slope of this line gives A.

Also, to confirm that the power of n in $(eqn \ C.21)$ should be 0.5, a more general equation was solved. $(Eqn \ C.21)$ can be rewritten with a general exponent of n and then the optimum value of this can be found from analysis.

$$(eqn \ C.21) \qquad \frac{L}{X} = \left(1 + A \frac{CoV}{(n)^{B}}\right)$$

hence
$$\frac{L}{X} - 1 = A \frac{CoV}{(n)^{B}}$$

and

/Τ \

and
$$\ln\left(\frac{L}{X}-1\right) = \ln(A) + \ln(CoV) - B\ln(n)$$

 $\ln\left(\frac{L}{X}-1\right) = \ln(A) - B\ln(n)$ (eqn C.22)
In this case by plotting $\ln\left(\frac{L}{X}-1\right) = \ln(n)$, then the slope gives $-B$ and the

intercept $\ln(A)$. In this way the optimum index for *n* can be evaluated for each analysis method. The results invariably gave answers near 0.5, confirming the form of the (*eqn* 5.2)

C.3.3 Variation in CoV used in prescribed CoV analyses

For the analysis methods in which a prescribed CoV is used, the value chosen by a producer may not match the value of the production. The use of a CoV that is slightly different to the true value of the material will introduce further errors in the analysis.

It is possible to determine the effect of variability in the prescribed CoV value, and this was tested with a Monte Carlo simulation. Two levels of variability in CoV were chosen as described in Section 4.1.3. The analyses gave repeatable results.

C.3.4 Success of data simulation

A major key to the success of Monte Carlo simulation is to create distributions that correctly model the data.⁸ A number of different data sets were used to fit distributions for each of the modelled parameters.

A distribution is used to model a single property for a single grade in a single production run. It simulates the testing of every piece of timber in-grade in the run. There are no data sets to compare this with, but a number of other data sets were used in the study.

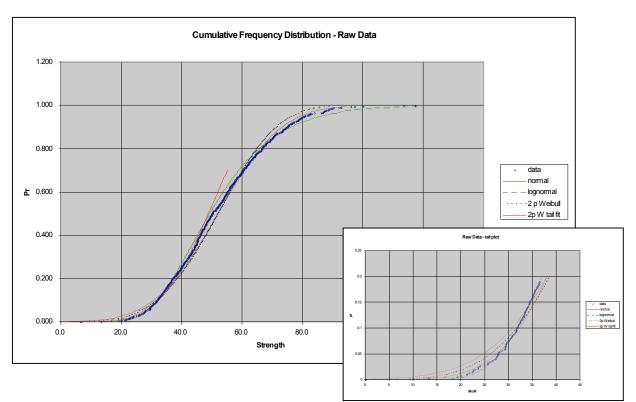


Figure C.2 – Data fit to bending strength data

C.3.4.1 Bending strength data

Figure C.2 shows some test bending strength data for a single mill and a single grade with a number of distributions fitted to it.

- The log-normal distribution fitted the data well over the complete range of the data.
- In the region of the lower tail, the log-normal distribution also followed the data well.

The log-normal distribution was used to model the strength of production. In determining the sensitivity of the analysis method to the assumption of the distribution of the production properties, the following distributions were also tried:

- Beta general
- Gamma
- Inverse Gaussian

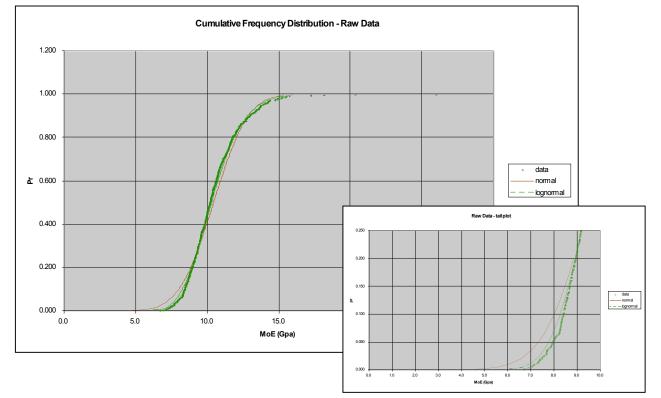
The results of the simulations did not differ with the different distributions tried. It was decided to use the log-normal distribution because it was understood already in the industry and it fitted the bending strength data as well as any of the others tried.

C.3.4.2 MoE data

Figure C.3 shows some MoE data for a single mill and a single grade with a number of distributions fitted to the data.

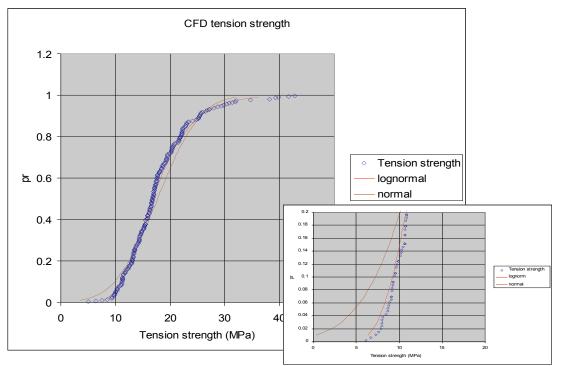
The MoE data fitted quite well to the log-normal distribution, and while the normal distribution give an accurate prediction of the average value and was well aligned

with the data near the top of the distribution, there was a misalignment between the normal distribution and the data at the tail.



The log-normal distribution was used to model the stiffness data from production runs.

Figure C.3 – Data fit to MoE data



C.3.4.3 Tension strength data

Figure C.4 – Data fit to tension strength data

Again, the tension strength data was well modelled by the log-normal distribution.

C.4 Results

C.4.1 Under and over-estimation of properties

Different methods of analysis only produce estimates of the property for the production over a given period, and can systematically underestimate or over-estimate the estimated property.

In general:

- In estimating the mean of a property, there is little systematic under-estimation or over-estimation.
- In estimating the 5% ile of a property, methods that use all test data overestimate the 5% ile.
- For larger numbers of specimens in a sample, there is lower magnitude of under- or over-estimation.
- For higher CoV of the property being measured, there is a higher magnitude of under- or over-estimation.

In this study, the following measure of over-estimation was used:

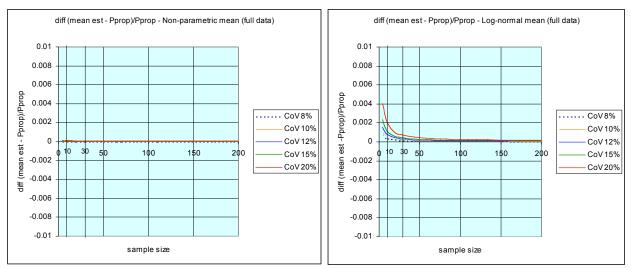
$$Over-estimation = \frac{(average value - production property)}{production property}$$
(eqn C.23)

This gives an estimate of the systematic over-estimation of the property as a fraction of the true value of the property. Note that this is not necessarily the situation in every case, as can be seen in the illustration in Table C.2. However, over a large number of estimations, a general trend can be seen for a given analysis method to under- or over-estimate the production parameter.

To facilitate comparison, all of the plots within a single Figure will have the same scale.

C.4.1.1 Estimation of the mean and 5%ile MoE

For modelling MoE, 1000 simulations were used. Statistical methods are quite robust in estimating the mean of a sample, and this was confirmed in the study.



(a) non-parametric estimation of mean Figure C.5 – Over

estimation of mean (b) estimation of mean using log-normal **Figure C.5 – Over-estimation of mean MoE**

Figure C.5 shows that:

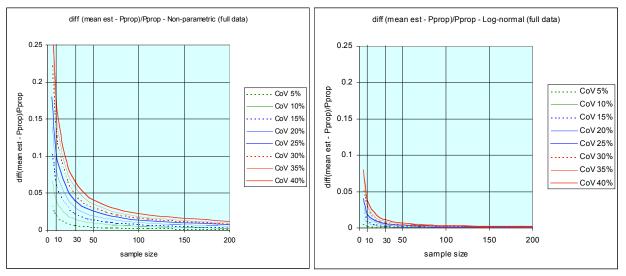
- There was no systematic error for any of the evaluations of mean MoE using the arithmetic mean of the test samples. (Figure C.5(a))
- However, where a log-normal distribution was fitted through the data, and the mean found from the log-normal distribution, there was a very small tendency to over-estimate the mean MoE from the data. Figure C.5(b) shows that this was generally less than 0.1% of the true mean MoE for 20 or more specimens per sample.

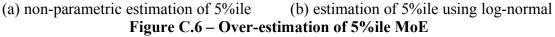
The systematic overestimation of the mean MoE by either analysis was negligible.

In contrast to estimations of the mean MoE, the study showed that both methods gave over-estimation of the 5% ile MoE. Figure C.6 shows that:

- For evaluations of 5%ile MoE using non-parametric methods on the test data, all data sets show an over-estimation of the 5%ile MoE. (Figure C.6(a)). The over-estimation is highest for small sample sizes and high CoV products. For sample sizes of 20 or more, normal CoV for MGP timber gives a systematic over-estimation of 5%ile MoE of 3% or less.
- Where a log-normal distribution was fitted through the data, and the 5%ile found from the fitted log-normal distribution, there was a much smaller tendency to over-estimate the 5%ile MoE from the data. Figure C.6(b) shows that this was generally less than 0.4% of the true 5%ile MoE for 20 or more specimens per sample. This is around 1/10th the systematic error of non-parametric estimations of 5%ile MoE

The systematic overestimation of the 5% ile MoE by either analysis was small, but much smaller for the 5% ile estimated from a log-normal distribution fitted through the data. Any tendency to over- or under-estimate production properties has been taken into account in the findings of the study.



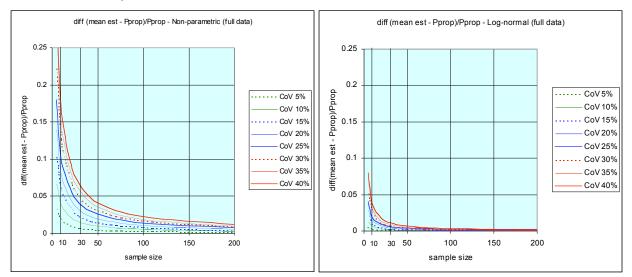


C.4.1.2 Estimation of 5%ile Strength from all test data

The key parameter for estimation of strength is the 5% ile strength. As there was more variability in strength data compared with MoE data, 2500 simulations were used to determine the distribution of the estimation of strength.

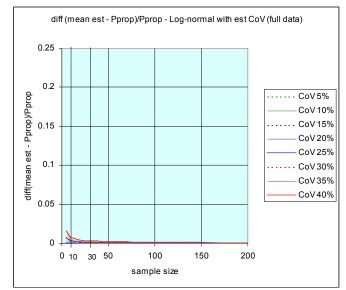
Figure C.7 shows the systematic overestimation of the 5% ile strength from analysis methods that use all of the data from the test series. (All test specimens are loaded to failure, and all failure loads are used in the analysis.) Figure C.7 shows that:

- The 5% ile is generally over-estimated by all analysis methods that use all of the test data.
- The systematic over-estimation of the non-parametric methods was significantly greater than for the methods based on a log-normal fit through the data.
- The overestimation is greater for higher coefficients of variation and for smaller sample sizes.
- The curves for the same CoV are identical in Figure C.6(a) and in Figure C.7(a) and also in Figure C.6(b) and Figure C.7(b). The overestimation is a function of the analysis method (5%ile by non-parametric methods or log-normal fit) and not the property being measured (5%ile strength or 5%ile MoE).



(a) non-parametric estimation of 5%ile

(b) estimation of 5%ile using log-normal



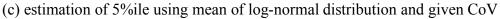


Figure C.7 – Over-estimation of 5%ile strength using full data set

Figure C.7 shows that the overestimation of the 5% ile for sample sizes of 20 or greater is less than 1 or 2 % for the log-normal fit methods.

The over-estimation of a 5% ile where the full data set is used has a logical explanation. Particularly where small sample sizes are used, it is likely that the tails of the distribution will not be represented in every sample. The mean value of the sample will be close to the production mean, but the outer extremities of the sample will be closer to the mean than it is for the complete production. It is therefore likely that all percentiles (including the 5% ile) calculated from the sample will be closer to the mean than those of the production.

The implication of this finding is that where small samples are taken, the general trend is for an over-estimation – giving the appearance that the production properties are quite satisfactory. However when larger samples are taken (even by amalgamating all of the smaller data sets), the over-estimation is smaller, making it appear that the production did not have the same margin of safety that the test results from the succession of smaller samples indicated.

C.4.1.3 Estimation of 5%ile strength from tail data

In contrast to the systematic over-estimation of 5% ile strength by analyses that used the whole of the data, Figure C.8 shows that there is a systematic under-estimation of the 5% ile where the distribution is fitted only to the tail for small sample sizes, but that for larger sample sizes, the property is often over-estimated.

Figure C.8 shows that:

- For small sample sizes, the under-estimation is very significant at up to 10%. (The sample size is the number of tested pieces, whether or not they failed or were included in the tail.)
- Typically, for sample sizes of 50 or more, there is a small over-estimation (much less than 5%).

The under-estimation for small sized samples is because with a small total sample size, there is a small probability of having test values in the lower tail. In many cases, the 5% ile cannot be estimated at all from the tail, because there were not enough data points. However, in some simulations of small sample sizes, the sampling program had thrown up a number of very low points, and an evaluation of the 5% ile was possible.

In this way the data was biased – no value if the sampling gave generally higher strength pieces, and a value if the sampling gave generally lower strength pieces. This effect tended to disappear for sample sizes of 30 to 50. For samples sizes of 50 or more, the true character of the distribution could be seen:

- Fitting data to log-normal distributions tended to give little or no overestimation for very large samples sizes in a similar trend to the full data set discussed in Section C.4.1.2.
- Fitting data to 2 parameter Weibull distributions gave a consistent but small over-estimation due to the shape of the curve at the tail.

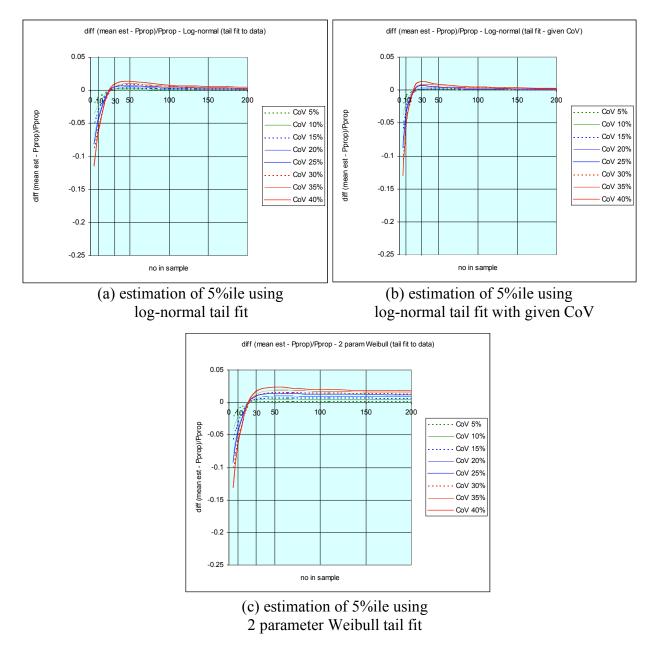


Figure C.8 – Over-estimation of 5%ile strength using tail data

C.4.2 Confidence in curves fitted to the results of the simulations

Each of the sets of simulations returned a confidence limit for the estimation of the production property. For each analysis method and each confidence level, there were a number of evaluations for:

- Different values of CoV for the production, and
- Different numbers of specimens in the sample.

An equation of the form given in (eqn C.24) was fitted to the analysis results:

Confidence Limit on Production property (L) =

$$L = X \left(1 + A \frac{CoV}{\sqrt{n}} \right)$$
 (eqn C.24)

(*eqn C.24*) was fitted to the confidence limits from the different simulations that made up each analysis method and each confidence level. For strength evaluations, there were 56 combinations of different CoV and sample size, and for stiffness evaluations, there were 35 combinations.

The confidence limits from the simulations could be compared with the confidence limits predicted by the equation, and a correlation coefficient (r^2) found. In (eqn C.24) \sqrt{n} can be written as $(n)^{0.5}$, with the power of n as 0.5. However, it was possible to see whether a different power of n gave a better fit to the data. For each set of results, the optimum power of n was found.

Table	e C.3 – Curv	e fitting to s	simulation r	esults	
5%ile strength from non	-parametric a	inalysis			
Confidence level	0.95	0.9	0.85	0.8	0.75
Optimum exponent of <i>n</i>	0.484	0.506	0.526	0.545	0.5698
A	-3.69789	-3.07219	-2.65129	-2.30915	-2.02086
r^2	0.991	0.998	0.999	0.998	0.996
	-3.698	-3.072	-2.651	-2.309	-2.021
5%ile strength from log-	normal analy	vsis			
Confidence level	0.95	0.9	0.85	0.8	0.75
Optimum exponent of <i>n</i>	0.536	0.544	0.554	0.563	0.576
A	-2.65855	-2.10389	-1.73106	-1.43076	-1.17217
r ²	0.994	0.993	0.991	0.988	0.987
	-2.658	-2.104	-1.731	-1.431	-1.172
5%ile strength from log-	normal analy	sis using giv	ven CoV (pr	odn value +/	- 5%)
Confidence level	0.95	0.9	0.85	0.8	0.75
Optimum exponent of <i>n</i>	0.395	0.374	0.353	0.333	0.310
A	-2.16558	-1.80641	-1.56428	-1.37204	-1.20416
r ²	0.994	0.990	0.9851	0.978	0.967
	-2.166	-1.806	-1.564	-1.372	-1.204
Average MoE from non-					
Confidence level	0.95	0.9	0.85	0.8	0.75
Optimum exponent of <i>n</i>	0.497	0.499	0.500	0.500	0.504
A	-1.64111	-1.28298	-1.03858	-0.84749	-0.6783
r ²	0.999	0.998	0.999	0.996	0.995
	-1.641	-1.283	-1.038	-0.847	-0.678

Table C.3 shows typical results for fitting the equation to the data.

Table C.3 – Curve fitting to simulation results

Table C.3 shows that in each case, the optimum exponent of *n* is close to 0.5, for all of the fits except the 5% ile strength from the log-normal distribution with a given CoV (within \pm 5% of the Production value). In the case of the analysis, which showed a different optimum exponent, the result was not particularly sensitive to the actual value, with the results for an exponent of 0.5 also giving a high correlation between the equation and the simulation results. An exponent of 0.5 was used for all relationships presented in this report.

The analysis was performed on each model considered, and the high correlation coefficient confirmed the form of the equation. Any analysis method that gave an r^2 for this correlation of less than 0.65 was not included in the outcomes of this report. The analysis methods presented in the outputs had an r^2 of more than 0.95.

C.4.3 Validation of the Outputs

There were a few outputs for which confirmation was available:

- Where a mean property is estimated, the confidence limits can be found by standard statistical methods^{1,2,10,11,12}. It was possible to compare the values found using the simulations with the standard confidence limits.
- Where a 5% ile was sought, this was not possible, but there are a few cases where a confidence limit on 5% ile strength is given in Standards or other literature.

C.4.3.1 Validation of CL for mean values

The form of the equation for the Confidence Limit (eqn C.24), is the same as for determination of confidence limits for the mean value from a normal distribution. The values of A from the estimation of the mean of the MoE, (where the values for A from the study were essentially the same for non-parametric, normal and log-normal distribution analyses) can be compared with the appropriate z values from the normal distribution.

Table C.4 – Comparison of CL for mean						
Confidence level	0.95	0.9	0.85	0.8	0.75	
Normal distribution (z)	-1.645	-1.282	-1.036	-0.842	-0.674	
Study (A)	-1.649	-1.290	-1.045	-0.854	-0.686	

Table C.4 – Comparison of CL for mean

Table C.4 shows that the methodology used in the simulations and subsequent analyses produced values for the confidence limit relationships that compared well with accepted statistical values for estimation of the mean.

In addition AS4063¹⁰ presents an equation for finding the 75% confidence limit of average MoE as:

$$CL_{0.75} = \overline{E} \left(1 - 0.7 \frac{CoV}{\sqrt{n}} \right)$$

In this case, the -0.7 is essentially the same value as the -0.69 in the study result. The Standard appears to use accepted statistical confidence calculations for MoE.

C.4.3.2 Validation of CL for 5%iles

There are no accepted statistical values for confidence limits on the estimation of the 5% ile from various distributions. However, there are two confidence estimates that have been accepted in timber testing, and they can be compared with the results of this study:

• US standards present the 75% CL on estimates of 5% ile strength of data obtained from tail fitting of a 2 parameter Weibull distribution as

$$CL_{0.75} = x_{0.05} \left(1 - 2.7 \frac{CoV}{\sqrt{n}} \right)$$
 with the CoV as the tail CoV from the

Weibull distribution and *n* as the total number of tests (not just the tail).

• European standards present the 75% CL on estimates of 5% ile from a lognormal analysis of the full data as

$$CL_{0.75} = x_{0.05} \left(1 - 1.25 \frac{CoV}{\sqrt{n}} \right)$$
 with the CoV from the log-normal

distribution of the full data set.

A comparison can be made with the results of the simulations and analyses used in this study. This is shown in Table C.5.

In this case, there were no adjustments made for under-or over-estimation of the production properties. This differs from the values calculated for the remainder of the report in which the confidence limit was compared with the actual production value rather than the mean estimate. Table C.5 shows data that was compatible with the data believed to be the foundation of the overseas analyses.

Table C.5 Comparison of CL for 570he strenge					
Confidence level	0.75	0.75			
Origin	US ⁷	Europe ¹⁷			
	2 parameter				
Analysis	Weibull	Log-normal			
CoV	Tail	full			
N	Full	full			
Adjustment for over-					
estimation	No	no			
Value in Standards (A)	~-2.7	-1.25			
Study (A)	-2.281	-1.044			
Ratio (Standards/Study)	1.18	1.19			

 Table C.5 – Comparison of CL for 5% ile strength

The standards seem to have a consistent factor of around 1.18 over the study results. Some conservatism in addressing the results of strength of timber has been built into the Standards.

The outcomes of this study did not incorporate any additional conservatism in the confidence limits presented for strength.

C.4.4 CoV for tail fitting 2 parameter Weibull distributions

It is well acknowledged that the CoV from a tail fit to a 2 parameter Weibull distribution gives a lower CoV than the CoV for the whole distribution⁷. Previous studies have indicated that it is about half of that for the whole distribution.

The study established a relationship between the CoV of the whole of the data and the appropriate CoV for the 2 parameter Weibull fit to the tail. This enabled comparisons of A values for (*eqn C.24*) between 2 parameter Weibull tail fits (where the CoV is less than for other methods).

The results of the simulations gave both the CoV for the whole of the data and the CoV determined from the 2 parameter Weibull fit through the tail of the data. A unique relationship was established that was independent of the CoV of the timber

being sampled. The ratio varied with the proof stress (or the maximum stress used to define the tail of the distribution).

Figure C.9 shows the relationship between the two CoVs as a multiplier plotted against the probability of the proof load – pr(proof). (For example, if the proof load is set at the 15% ile strength, then pr(proof) is 15%.) The data from the study is shown as points, and the thick black line is the line-of-best-fit through all of the test data.

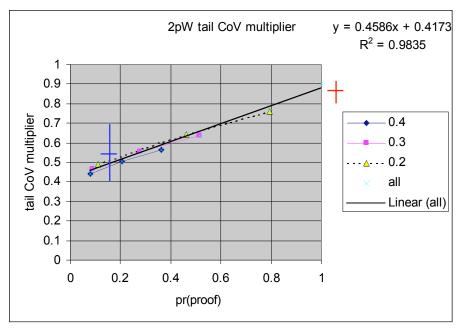


Figure C.9 Tail multiplier for 2 parameter Weibull tail CoV

The 2 parameter Weibull tail CoV is found by multiplying the CoV for the whole of the data by a multiplier given by (*eqn* C.24). This expression is a function of the probability of the proof stress in the whole of data distribution.

$$Mult_{CoV} = 0.4586 \, pr(proof) + 0.4173$$
 (eqn C.25)

For this equation, the following key points result:

- Taking the tail as the lower 15% of the data, gives the value of the multiplier as 0.486, or very close to 0.5.
- Taking the whole of the data (or fitting to 100% of the data), gives the value of the multiplier as 0.876 or very close to 0.9.

The analysis of the 2003 data for all radiata pine mills MGP10 gave the following values:

- Multiplier for a tail to 15% is 0.58, with values for individual mills ranging from 0.39 to 0.7. This data is shown as a blue cross in Figure 13.
- Multiplier for 2 parameter Weibull fit to the entire data is 0.88, with values for individual mills ranging from 0.85 to 0.93. This data is shown as a red cross in Figure 13.

This finding can be used to estimate the CoV that should be used in fitting a 2 parameter Weibull distribution to the tail of some data, knowing that the CoV for the product normally has a given value.

C.4.5 Proof testing

A number of mills make use of proof testing as part of their monitoring programs. (This is different from the proof testing provisions of AS/NZS4490 – used in Periodic Monitoring.) In proof testing for daily testing, while each sampled piece is loaded, only the very weakest pieces are broken. When the test reaches a stress at which it is clear that the specimen will not have a strength in the tail of the distribution, the test is stopped. This stress is known as the proof stress, and it is generally just a little higher than the expected 5% ile of the timber. Testing in this way produces MoE data for each test specimen, but a reduced number of strength data. There is no strength data for pieces where the proof stress was attained without a failure and the test was stopped.

In these cases, it is only possible to fit a distribution through the tail of the data represented by the results of failures. (All of the strengths in the tail fit are less than the proof stress.)

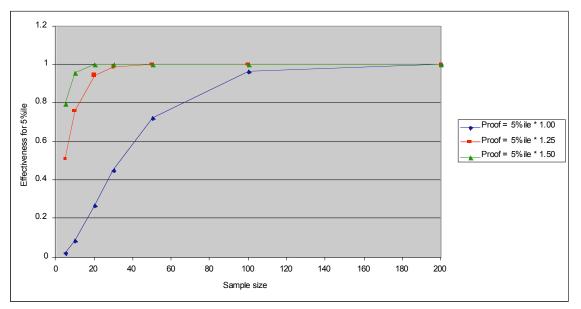


Figure C.10 – Effectiveness of determining 5%ile from tail fits with different proof loads (log-normal fit averaged over different CoVs)

The estimate of the 5% ile strength of the test samples is affected by the level of proof stress used:

- Where only the tail is used to fit a distribution, there are not many data points to enable a good fit to the data. While the fit may match the character of the tail well, if the tail is affected by one or two samples, then the fit may not represent the rest of the distribution particularly well. The lower the proof stress, the fewer points there are in the data set.
- Where the proof stress is close to the 5% ile value, there may not be enough failure results from small test samples to enable calculation of the 5% ile value.

Figure C.10 shows a plot of the effectiveness of estimating the 5% ile strength from tail data using different proof stresses. The value plotted is the fraction of successful evaluations when the proof stress is the given ratio of the actual 5% ile of the data. It shows that:

- The proof stress needs to be at least 1.25 of the actual 5% ile to have reasonable success.
- For proof testing, a sample size of at least 20 is required to give at least a 90% effectiveness of evaluations.

Figure C.10 was presented for a log-normal fit to the tail of the distribution, and the results for the 2 parameter Weibull fit showed a similar trend, but with lower effectiveness for small sample sizes.

The effectiveness of the determination was also a function of the CoV of the material. The method was more effective where the CoV was smaller. This is illustrated in Figure C.11 which shows similar information for a proof load of 1.25 times the 5%ile but plotted for different CoV materials.

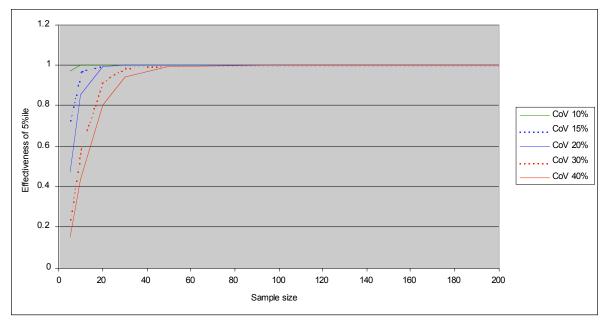


Figure C.11 – Effectiveness of determining 5%ile from tail fits on material with different CoVs (log-normal fit for a proof load of 1.25 times 5%ile)

These plots confirm the fact that tail fitting may have very low reliability with small data sets. A minimum size is 50 tests, with 100 more reliable.

In the data presented for 2 parameter Weibull tail fitting, the 15% ile stress is chosen as the proof stress.

Table C.6 shows the variation in A (from eqn C.24) with different levels of proof load. It shows that the values decrease as the proof stress increases. This indicates that there is less sampling error in the results for higher proof stresses.

Table C.6 – error from larger proof stress (Indicated by A in eqn C.24)

Confidence level	0.95	0.9
Proof stress = 5%ile * 1.00	-4.06	-2.87
Proof stress = 5%ile * 1.25	-3.03	-2.40
Proof stress = 5%ile * 1.50	-2.98	-2.38

C.4.6 CUSUM simulation

Simulations of a single CUSUM process were performed in order to demonstrate the confidence producers could reasonably have in its ability to respond to problems in products.

CUSUM is a method for giving a binary (pass/fail) result for properties of any production process. There are international standards that detail it. However, as a binary process, it is not suited to the analysis methods presented in this project. It is used in the timber industry in Australia and elsewhere, so a simple comparison with the methods outlined in this project is given in this section.

CUSUM is one of a number of statistically based Quality Control tools. It is a QC method that has been widely used in the Northern hemisphere and by a number of Australian and New Zealand mills. It is used because of the following reasons:

- It has been codified for structural timber (overseas)¹³
- It uses relatively small sample numbers to give a binary result
- It requires very little statistical knowledge from managers of timber grading operations
- It is an off-the shelf QC method.

C.4.6.1 CUSUM for MoE

The simulation of CUSUM for MoE involved generation of five random test samples from data with a known average MoE and a given CoV. For each simulation, the data was subjected to the standard CUSUM for variables procedure¹⁴:

- The average MoE of the test specimens was found (E_{avg}) .
- A SUM was evaluated from the previous CUSUM score, E_{avg} and the CUSUM variables.
- The CUSUM algorithm was used to evaluate the new CUSUM score.
- The CUSUM score enabled an interpretation as to whether or not the process was 'under control'.

The results of the analysis targeting MGP12 gave plots of the type shown in Figure C.12. Here the probability of the production being declared 'under control' is plotted against the average MoE of the population used to draw the samples. 95% probability means that there is a 95% chance of the production being declared 'under control'.

Figure C.12 shows that if the production MoE drops appreciably, the CUSUM system will reject it quite quickly.

- Where the production average MoE has fallen to a value less than 8.5 GPa (67% of the design value), then there is a probability of less than 96% that it will be classed as 'under control' by the CUSUM on the first shift after the change in properties. This massive reduction in properties has a 0.04% chance of passing on the second shift, so would almost certainly be caught on the second shift after the drop.
- The red line in Figure C.12 shows the performance of material in which the MoE has fallen to 11 GPa (87% of the design value). Here the first shift result has a 100% chance of being classes as 'under control', but after the second shift from the fall in properties, the CUSUM has 96% of being classed as

'under control'. By the third shift the probability is 66%, and after 10 shifts it is less than 0.1%.

• The blue line in Figure 16 is set at an average shift MoE of 12 MPa (95% of the design value). In this case, the probability of the first three shifts being classed as 'under control is more than 99%. It is only after 5 shifts that there is less than 95% confidence in having the production accepted.

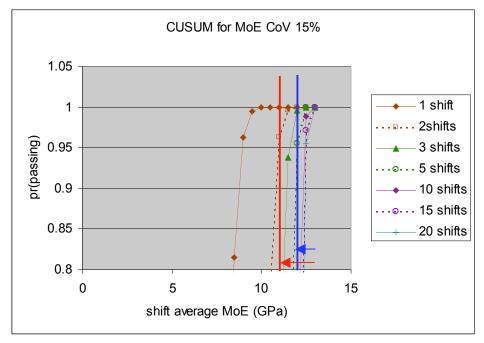


Figure C.12 – probabilities of MoE CUSUM accepting production

Table C.7 presents the 95% acceptance level for the production being classed by CUSUM as 'under control' for MGP12. This table gives the level that the average MoE for the shift has to drop to in order to have only 95% confidence of gaining an 'under control' outcome from the CUSUM after the designated number of shifts.

• For example for a CoV of 15%, after 2 shifts, the CUSUM result has a 95% chance of being 'under control' where the average MoE of the shift is 10.8 GPa or more.

	CoV	1 shift	2shifts	3 shifts	5 shifts	10 shifts	15 shifts	20 shifts
	20	9.2	11.3	11.9	12.3	12.7	12.8	12.8
	15	8.9	10.8	11.7	12	12.3	12.4	12.5
	12	8.7	10.7	11.4	11.9	12.3	12.4	12.4
	10	8.6	10.7	11.3	11.9	12.3	12.4	12.4
	8	8.4	10.7	11.3	11.8	12.1	12.2	12.3

Table C.7 – shift average MoE with 95% probability of being 'under control'

Table C.7 shows that as long as the MoE is above the design value on a sustained basis, there is little chance of a rejection of product. On the other hand, shifts that have dropped to less than 10 MPa do not have 95% confidence in being passed after the second consecutive shift at that level.

C.4.6.2 CUSUM for strength

The simulation of CUSUM for strength involved generation of five random test samples from data with a known 5% ile strength and a given CoV. For each simulation, the data was subjected to the standard CUSUM for attributes procedure¹⁴:

- The number of failures at a strength less than the threshold strength were counted (*N*).
- A SUM was evaluated from the previous CUSUM score, *N* and the CUSUM variables.
- The CUSUM algorithm was used to evaluate the new CUSUM score. The algorithm differed slightly depending on whether the previous CUSUM was 'under control' or not.
- The CUSUM score enabled an interpretation as to whether or not the process was 'under control'.

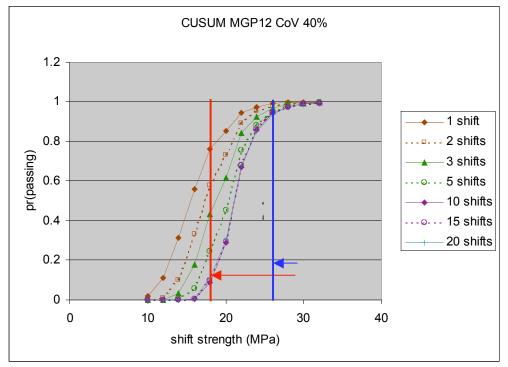


Figure C.13 – probabilities of strength CUSUM accepting production

The results of the analysis targeting MGP12 properties gave plots of the type shown in Figure C.13. Here the probability of the production being declared 'under control' is plotted against the 5% ile strength of the population used to draw the samples. 95% probability means that there is a 95% chance of the production being declared 'under control'.

Figure C.13 shows that if the shift strength is very low, CUSUM will reject it quite quickly, but if the shift strength is quite close to the design value, it takes much longer to be rejected.

- A shift 5% ile strength to 12 MPa has only an 11% chance of being classed as 'under control' in its first shift. (This is equivalent to saying that it has an 89% chance of being rejected.
- The red line in Figure C.13 illustrates the scenario of the product 5%ile strength dropping suddenly to 18 MPa (64% of the design value). Under these

circumstances, there is a 76% chance of passing the first shift. However, the probability of passing the second shift after the drop is less (58%) and lower again after three shifts (44%). The probability of it still being classed as 'under control' by CUSUM after 20 shifts is 11%.

• The blue line in Figure C.13 illustrates the scenario of the product 5%ile strength being maintained at 24 MPa (86% of the design value). In this case, there is a 97.5% chance of the production being classed as 'under control' in the first shift at the new value. (It only has s 2.5% chance of being caught.) The probability of passing the production decreases slightly with each continuing shift, but after 10 shifts, it has stabilised at 85%. It has a relatively low probability of rejection.

As the shift strength gets closer to the design strength (in this case, 28 MPa) it is much more likely to be accepted by the CUSUM criteria¹⁵. Table C.8 presents the shift strength for production that has a 95% probability of being accepted.

	sinne	shine serengen witch >0 /0			prosubling of sening under control			
CoV	1 shift	2 shifts	3 shifts	5 shifts	10 shifts	15 shifts	20 shifts	
40	22	24	25	26	27	27	27	
35	23	24	25	26	27	27	27	
30	23	24	25	25	26	26	26	
25	23	24	25	25	25	26	26	

Table C.8 – shift strength with 95% probability of being 'under control'

Table C.8 shows that there is a high probability of acceptance of material for a number of shifts where the strength is a little less than the design strength, but not appreciably lower than the design value.

Table C.8 does not appear to be very sensitive to the CoV of the production material. At a probability level of 50%, there is a little more sensitivity, but the process does not vary much over the range of CoVs expected in sawn timber.

C.4.6.3 Comparison between CUSUM for strength and MoE

The red line in Figure C.12 (87% of the design MoE) can be compared with the blue line in Figure C.13 (86% of the design strength).

- For the MoE (the red line in Figure C.12) the first shift result has a 100% chance of being classed as 'under control', but after the third shift the probability is 66%, and after 10 shifts it is less than 0.1%.
- For the strength (the blue line in Figure C.13) the first shift result has a 97.5% chance of being classed as 'under control', but after 10 shifts, it has stabilised at 85%. It has a relatively low probability of rejection.

A fall in MoE value has a much higher chance of being caught compared with a fall of the same relative magnitude in strength. In both cases, there is a small chance of detection in the first shift unless the fall in properties is of the order of one grade.

The outputs from CUSUM from the same examples can also be compared with outputs from other methods of analysing the same data. For MoE:

• For the case of an evaluation at the end of a single shift – using only five data points.

- Using CUSUM, one shift at 11 GPa (87% of the design MoE) gave a 100% chance of being classed as good, or a 0% chance of being detected as a problem.
- Had the non-parametric average MoE been found from the same 5 samples, then the result would have had a 1% chance of being greater than 12.7 GPa clearly indicating a problem. Another way of considering it, is that for a 90% Confidence Limit, the Test Comparison Value would be 13.91 GPa and the student t test gives a 0.8% probability of a sample of 5 drawn from a production with a mean of 11 GPa and a CoV of 15% exceeding that value.
- For the case of an evaluation at the end of two production shifts after which a total of 10 data points are available and a decision will be made on the acceptance of both shifts taken as a group.
 - Using CUSUM, two shifts at 11 GPa (87% of the design MoE) gave a 96% chance of being classed as good, (based on a total of 10 pieces tested) or a 4% chance of being detected as a problem.
 - Had the non-parametric average MoE been found from the same 10 samples, then the result would have had a 0.1% chance of being greater than 12.7 GPa clearly indicating a problem. Another way of considering it, is that for a 90% Confidence Limit, the Test Comparison Value would be 13.53 GPa and the student t test gives a 0.05% probability of a sample of 10 drawn from a production with a mean of 11 GPa and a CoV of 15% exceeding that value.

Therefore, using the same data to find the non-parametric average gives better certainty in detecting problem batches compared with the CUSUM method. (In this case 95% confidence for the raw data compared with 4% after two shifts using CUSUM.)

A similar comparison can be made between the strength from CUSUM and the strength from a log-normal fit through the same data.

- Based on the results of a single shift:
 - Using CUSUM, a single shift at 24 MPa gave a 97.5% chance of returning an 'under control' result. This equates to a 2.5% chance of detecting a problem at the end of that shift in which 5 samples were tested.
 - Had a log-normal fit to the five points been used to estimate a 5%ile with a CoV of 35%, then there is a 27.1% chance of the 5%ile being greater than the 28 MPa design value, based on a population with a 5%ile strength of 24 MPa. The probability of exceeding a Test Comparison Value (for any Confidence Level) will be much less than this. There is clearly a problem indicated after one shift.
- Based on the results of 10 shifts. (Assuming that the analysis is performed after 10 shifts and a decision made on the quality of that 10 shifts as a block)
 - Using CUSUM, ten shifts at 24 MPa gave an 85% chance of returning an 'under control' result based on a total of 50 test specimens. This equates to a 15% chance of detecting a problem.
 - Had a log-normal fit to the fifty points been used to estimate a 5% ile with a CoV of 35%, then there is a 1.6% chance of the 5% ile being greater than the 28 MPa design value, based on a population with a

5% ile strength of 24 MPa. The probability of exceeding a Test Comparison Value (for any Confidence Level) will be much less than this. It is highly unlikely that the problem would remain undetected after 10 shifts using this method of analysis.

For both strength and MoE, the analysis of the data to estimate properties directly gives a more reliable outcome than the use of CUSUM on the same data in comparing both the results of a single shift and the results of a number of shifts taken together.

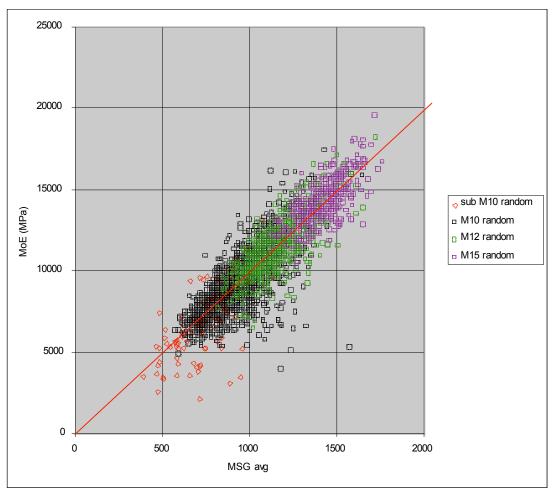
C.4.7 Finding mean MoE from MSG data

There is a high correlation between mean MoE determined from random position testing on random samples and the mean of the MSG averaged over the length of the same pieces, provided a sufficient number of pieces is used.

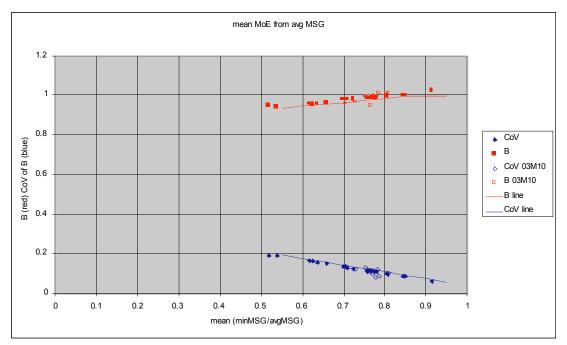
Figure C.14 shows a scatter plot of random position MoE and MSG averaged over the length of the same pieces. The data has been sourced from the 2003 MGP study for only 90×35 timber and covers a range of grades. It also shows the equality line. The data appears to be quite close to the equality line, though there is a scatter around the line. However, different relationships may exist for material from different resources.

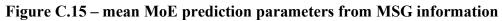
It was possible to determine the ratio (mean MoE/avg MSG) for different groupings of the data. A very strong correlation was found between the character of the timber and the ratio. Figure C.15 shows a plot taken from the same data shown in Figure C.14 in which the ratio is plotted against (min MSG/avg MSG) for the same grouping of timber.

- Figure C.15 shows red squares and lines for the ratio (mean MoE / average MSG). The full red boxes are from the grouped data shown in Figure C.14, and the data plots in a straight line. As well, each of the mills in the 2003 study were used to derive points for their MGP10 grades and these are shown as open boxes. The linear relationship (*eqn C.26*) is shown as the red line.
- Figure C.15 shows blue diamonds for the calculated value of CoV for the grouped data that gave the red boxes in the same plot. The CoV of the ratio also is very linear. Again the CoV was evaluated for the ratio for MGP10 from each of the mills in the 2003 study, and plotted as open blue diamonds. Again, the mill data falls very close to the linear approximation. The linear relationship (*eqn C.27*) is shown as the blue line.









The linear relationships plotted in Figure C.15 are given as (eqn C.26) and (eqn C.27)

$$\left(\frac{meanMoE}{avgMSG}\right) = B = 0.827 + 0.197 \left(\frac{\min MSG}{avgMSG}\right)$$

$$(eqn \ C.26)$$

$$CoV_{\left(\frac{meanMoE}{avgMSG}\right)} = 0.377 - 0.334 \left(\frac{\min MSG}{avgMSG}\right)$$

$$(eqn \ C.27)$$

The prediction of mean MoE is therefore: meanMoE = B avgMSG with B given in (eqn C.26)

The shift mean MoE can be estimated from the average MSG for the shift multiplied by *B* given by (*eqn C.26*). *B* is a function of the character of the product and it can be found from the average (min MSG) for the grade and the average (avg MSG) for the grade.

In order to estimate the confidence of the data, the CoV of the ratio is needed. It can also be found from the average (min MSG) for the grade and the average (avg MSG) for the grade. This gives rise to a slightly different form of the equation for the confidence limit as shown in ($eqn \ C.28$). When comparing it with ($eqn \ C.24$) it can be seen that it has the factor B involved.

CL on Mean MoE =
$$CL_E = \overline{MSG} B\left(1 + A \frac{CoV}{\sqrt{n}}\right)$$
 (eqn C.28)

Here the CoV is the CoV of *B* and is given by (*eqn C.26*) and *n* is the number of pieces of wood that contribute to the average.